

Remark 1: It is possible to prove the following stronger result (useful in particular if $|k|$ is very large):

$$\sup_{k \in \mathbb{Z}} \max\left(1, \left(\frac{k-an}{\sqrt{n}}\right)^2\right) \left| \sigma\sqrt{n} \mathbb{P}(Z_n=k) - \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} \left(\frac{k-an}{\sigma\sqrt{n}}\right)^2\right) \right| \xrightarrow{n \rightarrow \infty} 0$$

See Principles of Random Walk, Chapter 7, by Spitzer

Remark 2: When $\sigma^2 = \infty$, it is possible to extend the Local Limit Theorem when the law of Z_1 belongs to the domain of attraction of a stable law of index $\alpha \in (0, 2)$

If $\mathbb{P}(Z_1 < -1) = 0$, this is equivalent to the fact that $\mathbb{P}(Z_1 \geq k) = \frac{L(k)}{k^\alpha}$ where L satisfies

$$\frac{L(ka)}{L(n)} \xrightarrow{n \rightarrow \infty} 1$$

for every $a > 0$

6) Applications of the Local Limit Theorem

Recall that $W_n = X_1 + \dots + X_n$ with $\mathbb{P}(X_1 = i) = \mu(i)$, $i \geq -1$ and that $\mathbb{E}[X_1] = \sum_{i \geq -1} i \mu(i) = \sum_{i \geq 0} i \mu(i) = -1$

Assume that μ is critical and has finite variance σ^2 .
 That $\mathbb{E}[X_1] = 0$, $\mathbb{E}[X_1^2] = \sum_{i \geq -1} i^2 \mu(i) = \sum_{i \geq 0} (i-1)^2 \mu(i) = \sum_{i \geq 0} i^2 \mu(i) - 1 = \sigma^2$

Definition: We say that μ is aperiodic if the maximal integer $h \geq 1$ such that $\sum_{i \geq 1; \mu(i) > 0} \mathbb{Z} \subset h\mathbb{Z}$ is $h=1$. Or, equivalently, that $\gcd \sum_{i \geq 1; \mu(i) > 0} \mathbb{Z} = 1$.

Lemma: Assume that μ is aperiodic and $\mu(0) > 0$. Then $\exists N$ s.t. $n \geq N \Rightarrow \mathbb{P}_\mu(|T|=n) > 0$

Proof: by aperiodicity, we can find i and j s.t. $\mu(i) > 0, \mu(j) > 0$ and $\gcd(i, j) = 1$.

Take $N = ij$.

Fix $n \geq N+1$. let $0 \leq b \leq i-1$ be the integer such that $bj \equiv n-1 \pmod{i}$ (possible since $\gcd(i, j) = 1$)

Hence there exists $a \in \mathbb{Z}$ such that $n-1 - bj = ai$.

But $bj < ij \leq n-1$. Hence $a \geq 0$.

Thus $n-1 = ai + bj$ with $a \geq 0, b \geq 0$.

Hence with positive probability a $\mathbb{C}W$ tree has a vertices with i children, b vertices with j children and all the other vertices 0 children. But such a tree has n vertices in total. Hence $\mathbb{P}_\mu(|T|=n) > 0$. \square

Remark: If μ aperiodic, then X_1 is aperiodic (in the sense defined in lecture 5).

(Indeed, if $\text{Support}(X_1) \subset c+h\mathbb{Z}$, then $\sum_{i \geq 1; \mu(i) > 0} \mathbb{Z} \subset c-1+h\mathbb{Z}$ and $\mu(i) > 0, \mu(j) > 0 \Rightarrow h | i-j$)

* Application of the Local Limit Theorem to the study of the size of $\mathbb{C}W$ trees

Assume that μ is critical, aperiodic, with finite variance.

We have seen that $\mathbb{P}_\mu(|T|=n) = \frac{1}{n} \mathbb{P}(W_n = -1)$.

By the local limit theorem, $\mathbb{P}(W_n = -1) = \frac{1}{\sqrt{2\pi\sigma^2 n}} e^{-\frac{1}{2} \left(\frac{-1}{\sigma\sqrt{n}}\right)^2} + \frac{\varepsilon(k, n)}{\sigma\sqrt{n}}$.

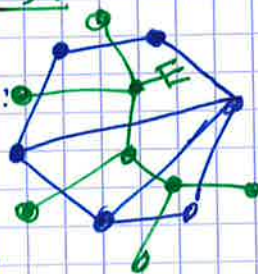
with $\sup_{k \in \mathbb{Z}} \varepsilon(k, n) \xrightarrow{n \rightarrow \infty} 0$. Hence $\mathbb{P}(W_n = -1) \sim \frac{1}{\sigma\sqrt{2\pi} \sqrt{n}}$.

Hence $\mathbb{P}_\mu(|T|=n) \sim \frac{1}{\sqrt{2\pi n^2}} \times \frac{1}{n^{3/2}}$

(in particular, note that $\mathbb{E}[|T|] = \infty$, which is consistent with the fact that $\mathbb{E}_\mu[|T|] = \mathbb{E}\left[\sum_{n \geq 0} z_n\right] = \sum_{n \geq 0} \mathbb{E}[z_n] = \sum_{n \geq 0} 1 = \infty$)

Application to counting dissections.

Recall the construction of a dissection of the dual tree $T(D)$:



If D_n is a uniform dissection of P_n , we have seen that the law of $T(D_n)$

is $\mathbb{P}_\mu(\cdot \mid \lambda(T) = n-1)$

where $\mu(0) = \frac{1-c}{1-c}$, $\mu(1) = 0$, $\mu(i) = c^{i-1}$ for $i \geq 2$.

We choose c so that μ is critical: $c = 1 - \frac{\sqrt{2}}{2}$

then $\mu(0) = 2 - \sqrt{2}$, $\mu(1) = 0$, $\mu(i) = \left(\frac{2 - \sqrt{2}}{2}\right)^{i-1}$ for $i \geq 2$.

Hence:

$\frac{1}{|D_n|} = \mathbb{P}_\mu(T = \text{tree with } n-1 \text{ leaves} \mid \lambda(T) = n-1) = \frac{\mu(n-1) \cdot \mu(0)}{\mathbb{P}_\mu(\lambda(T) = n-1)}$


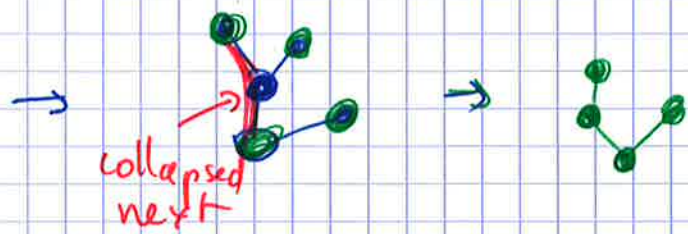
Hence $|D_n| = \frac{(2 - \sqrt{2})^3}{4} \cdot \frac{1}{(3 - 2\sqrt{2})^n} \cdot \mathbb{P}_\mu(\lambda(T) = n-1)$

How can we estimate $\mathbb{P}_\mu(\lambda(T) = n-1)$?

We use a construction due to Rizzolo (2011!) which maps a $\mathcal{G}W$ tree with n leaves to a $\mathcal{G}W$ tree with n vertices:

Start with a tree T with n leaves. Then, iteratively, "collapse" the edges of T that lead to the left-most leaf to get \tilde{T}

For example:

Rizzolo showed that if T is a $\mathbb{G}W_\mu$ tree, then \tilde{T} is a $\mathbb{G}W_{\tilde{\mu}}$ tree, and:

that μ critical with variance σ^2
 $\Rightarrow \tilde{\mu}$ critical with variance $\sigma^2/\mu(0)$.

Hence $\mathbb{P}_\mu(|T|=n-1)$
 $= \mathbb{P}_\mu(|\tilde{T}|=n-1)$
 $= \mathbb{P}_{\tilde{\mu}}(|T|=n-1) \sim \frac{1}{2\pi\sigma^2/\mu(0)} \times \frac{1}{n^{3/2}}$

We conclude by putting the pieces together that

$$|P_n| \sim \frac{1}{4\sqrt{\pi}} \cdot \sqrt{89\sqrt{2}-140} (3+2\sqrt{2}) n^{-3/2}$$

⊛ Number of children of the root.

We know that $\mathbb{P}_\mu(K_\emptyset = i) = \mu(i)$.

But what about $\mathbb{P}_\mu(K_\emptyset = i \mid |T|=n)$ as $n \rightarrow \infty$?

The coding by the Lukasiewicz path gives that

$$\mathbb{P}_\mu(K_\emptyset = i \mid |T|=n) = \mathbb{P}(X_{\pm} = i-1 \mid \sum_{\pm} = n)$$

$$= \frac{\mathbb{P}(X_{\pm} = i-1, \sum_{\pm} = n)}{\mathbb{P}(\sum_{\pm} = n)}$$

$$= \frac{1}{\mathbb{P}(\sum_{\pm} = n)} \cdot \mu(i) \cdot \mathbb{P}(\sum_{\pm} = n-1)$$

$$= i\mu(i) \times \frac{n}{n-1} \times \frac{\mathbb{P}(W_{n-1} = -i)}{\mathbb{P}(W_n = -1)}$$

since $\mathbb{P}(\sum_{\pm} = k) = \frac{c}{k} \mathbb{P}(W_{\pm} = -k)$

$\xrightarrow[n \rightarrow \infty]{} i\mu(i)$ by the local limit theorem.

Local limits of Galton-Watson trees.

Here μ is an offspring distribution on \mathbb{Z}_+ . We assume that μ is aperiodic.

1) Critical and finite variance case.

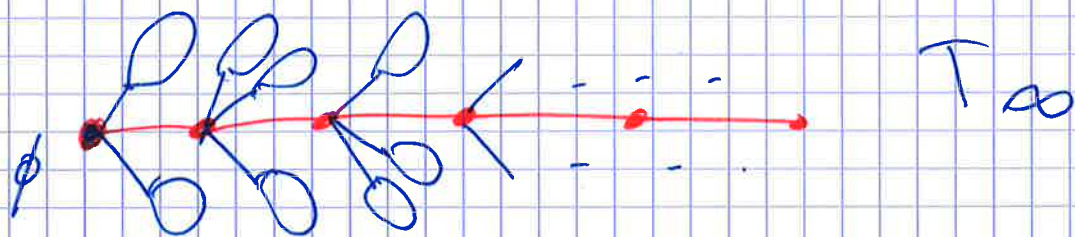
Assume that $\sum_{i \geq 0} i \mu(i) = 1$ and $\sigma^2 = \sum_{i \geq 0} i^2 \mu(i) - 1 \in (0, \infty)$.

Let T_n be a random tree with law $\mathbb{P}_\mu(\cdot \mid |T| = n)$.

Goal: T_n converges "locally" to an ∞ tree. T_∞

Define T_∞ as the following ∞ tree:

- there are 2 types of vertices: normal and mutant
- the number of children of "normal" vertices have law μ , and they are normal
- the number of children of mutant vertices have law $\bar{\mu}$ defined by $\bar{\mu}(i) = i \mu(i)$, and all are normal, except one which is mutant and which is chosen uniformly among the children:



For $t \in \mathbb{T}$, set $Z_k(t) = \{u \in t; |u| = k\}$
and $[t]_k = \{u \in t; |u| \leq k\}$

Prop

For every finite tree t_0

$$\mathbb{P}_\mu([T_\infty]_k = t_0) = Z_k(t_0) \mathbb{P}_\mu([T]_k = t_0)$$

Proof: exercise.

Assume that μ is critical, aperiodic, has finite variance

Thm

For every finite tree τ_0 and $k \geq 0$,

$$\mathbb{P}([\mathcal{T}_n]_k = \tau_0) \xrightarrow{n \rightarrow \infty} \mathbb{P}([\mathcal{T}_\infty]_k = \tau_0)$$

We say that \mathcal{T}_n converges locally in distribution to \mathcal{T}_∞ .

Remarks

- The local convergence may be defined by a metric
- If a distinguished point is not given, one can choose it uniformly at random, in which case one speaks of local weak convergence.

Proof: By the proposition, it is enough to check that for every tree τ_0 of height k we have

$$\mathbb{P}([\mathcal{T}_n]_k = \tau_0) \rightarrow Z_k(\tau_0) \mathbb{P}_\mu([\mathcal{T}]_k = \tau_0).$$

But

$$\mathbb{P}([\mathcal{T}_n]_k = \tau_0) = \mathbb{P}_\mu([\mathcal{T}]_k = \tau_0) \times \frac{\mathbb{P}_{\mu, Z_R(\tau_0)}(|\mathcal{R}| = n - |\tau_0|)}{\mathbb{P}_\mu(|\mathcal{T}| = n)}$$

$$= \mathbb{P}_\mu([\mathcal{T}]_k = \tau_0) \times \frac{Z_k(\tau_0)}{n - |\tau_0|} \times n \times \frac{\mathbb{P}(W_{n-|\tau_0|} = -\frac{Z_k(\tau_0)}{n})}{\mathbb{P}(W_n = -1)}$$

(Local limit theorem)

$\xrightarrow{n \rightarrow \infty}$

This completes the proof. \square

Remark: the theorem is true even if $\sigma^2 = \infty$, indeed it is known that $\forall m, k \in \mathbb{Z}$, (Strong ratio limit theorem for random walks)

$$\frac{\mathbb{P}(W_{n-m} = -k)}{\mathbb{P}(W_n = -1)} \xrightarrow{n \rightarrow \infty} \text{Theorem for random walks}$$