

*The Brownian triangulation:
a universal limit for random plane non-crossing
configurations*
(joint work with Nicolas Curien)

Igor Kortchemski (Université Paris-Sud, Orsay, France)

MIT Probability Seminar, April 2nd 2012

Motivations

Let $(X_n)_{n \geq 1}$ be a sequence of “discrete” objects converging towards a “continuous” object X :

$$X_n \xrightarrow[n \rightarrow \infty]{} X.$$

Motivations

Let $(X_n)_{n \geq 1}$ be a sequence of “discrete” objects converging towards a “continuous” object X :

$$X_n \xrightarrow[n \rightarrow \infty]{} X.$$

Several consequences:

- *From the discrete to the continuous world*: if a property \mathcal{P} is satisfied by all the X_n and passes to the limit, then X satisfies \mathcal{P} .

Motivations

Let $(X_n)_{n \geq 1}$ be a sequence of “discrete” objects converging towards a “continuous” object X :

$$X_n \xrightarrow[n \rightarrow \infty]{} X.$$

Several consequences:

- *From the discrete to the continuous world:* if a property \mathcal{P} is satisfied by all the X_n and passes to the limit, then X satisfies \mathcal{P} .
- *From the continuous world to the discrete world:* if a property \mathcal{P} is satisfied by X and passes to the limit, X_n satisfies “approximately” \mathcal{P} for n large.

Motivations

Let $(X_n)_{n \geq 1}$ be a sequence of “discrete” objects converging towards a “continuous” object X :

$$X_n \xrightarrow[n \rightarrow \infty]{} X.$$

Several consequences:

- *From the discrete to the continuous world:* if a property \mathcal{P} is satisfied by all the X_n and passes to the limit, then X satisfies \mathcal{P} .
- *From the continuous world to the discrete world:* if a property \mathcal{P} is satisfied by X and passes to the limit, X_n satisfies “approximately” \mathcal{P} for n large.
- *Universality:* if $(Y_n)_{n \geq 1}$ is another sequence of objects converging towards X , then X_n and Y_n share approximately the same properties for n large.

Motivations

Let $(X_n)_{n \geq 1}$ be a sequence of “discrete” objects converging towards a “continuous” object X :

$$X_n \xrightarrow[n \rightarrow \infty]{} X.$$

Several consequences:

- *From the discrete to the continuous world*: if a property \mathcal{P} is satisfied by all the X_n and passes to the limit, then X satisfies \mathcal{P} .
- *From the continuous world to the discrete world*: if a property \mathcal{P} is satisfied by X and passes to the limit, X_n satisfies “approximately” \mathcal{P} for n large.
- *Universality*: if $(Y_n)_{n \geq 1}$ is another sequence of objects converging towards X , then X_n and Y_n share approximately the same properties for n large.

What is the sense of the convergence when the objects are **random**?

→ **Convergence in distribution**

Convergence in distribution

Let $(X_n)_{n \geq 1}$ and X be random variables with values in a metric space (\mathcal{E}, d) . X_n **converges in distribution** towards X if

Convergence in distribution

Let $(X_n)_{n \geq 1}$ and X be random variables with values in a metric space (\mathcal{E}, d) . X_n **converges in distribution** towards X if

$$\mathbb{E} [F(X_n)] \xrightarrow[n \rightarrow \infty]{} \mathbb{E} [F(X)]$$

for every bounded continuous function $F : \mathcal{E} \rightarrow \mathbb{R}$.

Convergence in distribution

Let $(X_n)_{n \geq 1}$ and X be random variables with values in a metric space (\mathcal{E}, d) . X_n **converges in distribution** towards X if

$$\mathbb{E} [F(X_n)] \xrightarrow[n \rightarrow \infty]{} \mathbb{E} [F(X)]$$

for every bounded continuous function $F : \mathcal{E} \rightarrow \mathbb{R}$.

We write:

$$X_n \xrightarrow[n \rightarrow \infty]{(d)} X$$

Outline

I. THE DISCRETE OBJECT

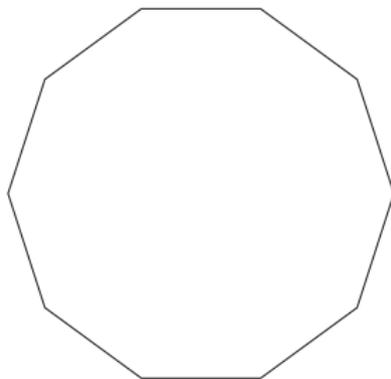
II. THE LIMITING CONTINUOUS OBJECT

III. PROVING THE CONVERGENCE

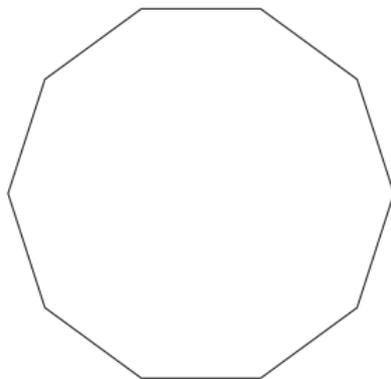
IV. APPLICATION TO THE STUDY OF UNIFORM DISSECTIONS

I. THE DISCRETE OBJECTS

Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).

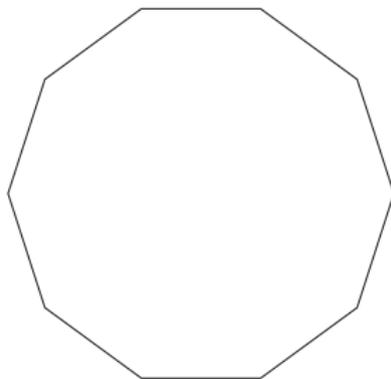


Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



Framework: choose a random **non-crossing** configuration obtained from the vertices of P_n , that is a collection of non-intersecting diagonals.

Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



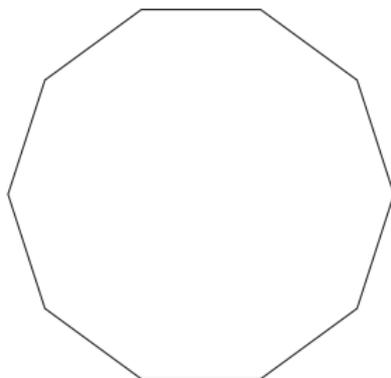
Framework: choose a random **non-crossing** configuration obtained from the vertices of P_n , that is a collection of non-intersecting diagonals.

What happens for n large?

Case of dissections of P_n .

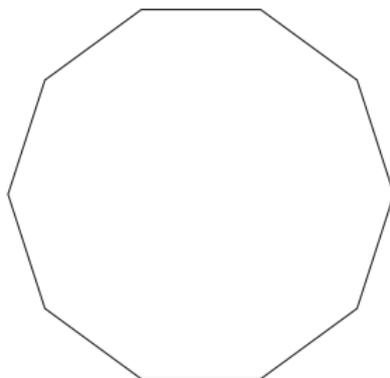
Dissections

Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



Dissections

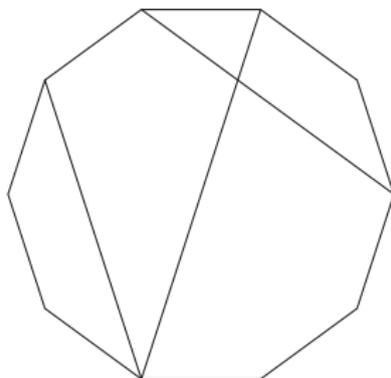
Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



A **dissection** of P_n is the union of the sides of P_n and of a collection of diagonals that may intersect only at their endpoints.

Dissections

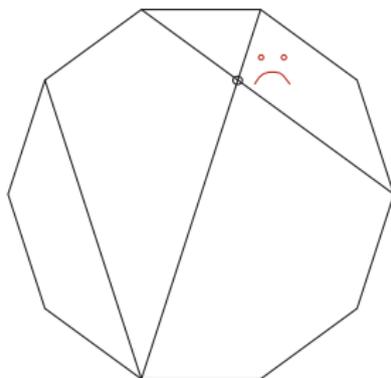
Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



A **dissection** of P_n is the union of the sides of P_n and of a collection of diagonals that may intersect only at their endpoints.

Dissections

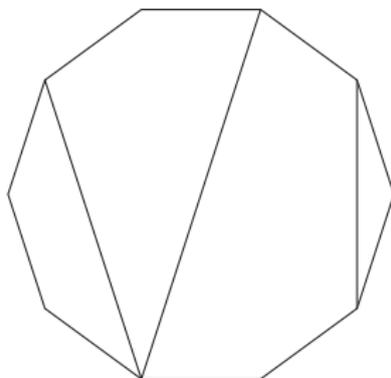
Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



A **dissection** of P_n is the union of the sides of P_n and of a collection of diagonals that may intersect only at their endpoints.

Dissections

Let P_n be the polygon whose vertices are $e^{\frac{2i\pi j}{n}}$ ($j = 0, 1, \dots, n-1$).



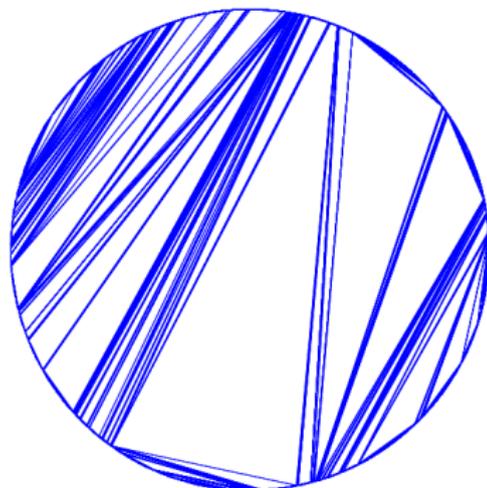
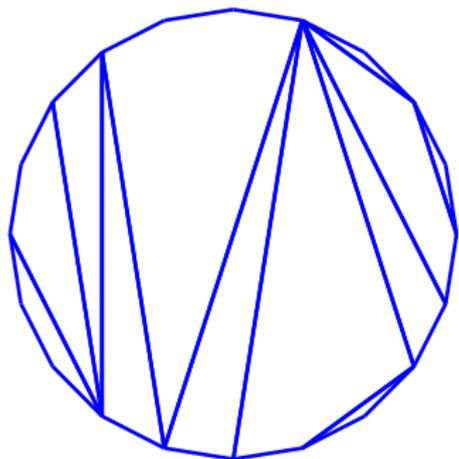
A **dissection** of P_n is the union of the sides of P_n and of a collection of diagonals that may intersect only at their endpoints.

Dissections

Let \mathcal{D}_n be a random dissection, chosen **uniformly** at random among all dissections of P_n . What does \mathcal{D}_n look like when n is large?

Dissections

Let \mathcal{D}_n be a random dissection, chosen **uniformly** at random among all dissections of P_n . What does \mathcal{D}_n look like when n is large?

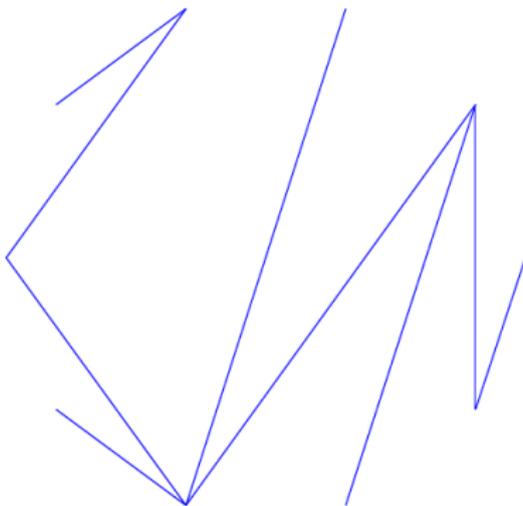


Samples of \mathcal{D}_{18} and \mathcal{D}_{15000} .

Case of **non-crossing** trees of P_n .

Non-crossing trees

Example of a **non-crossing** tree of P_{10} :

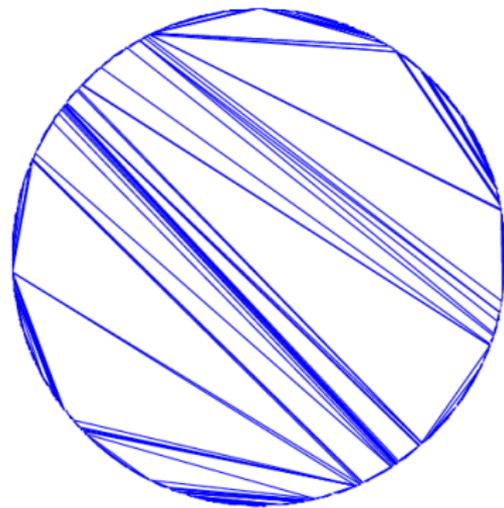
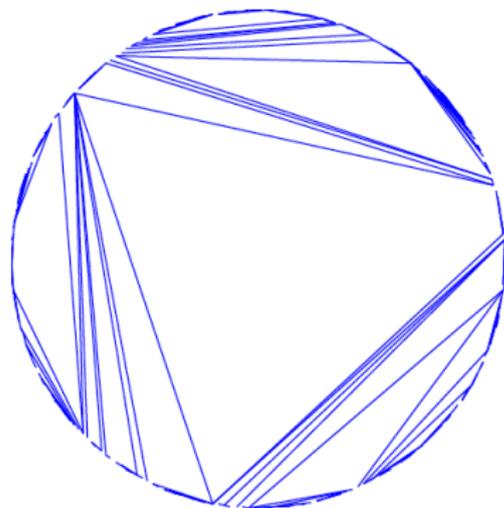


Non-crossing trees

Let \mathcal{T}_n be a random **non-crossing** tree, chosen **uniformly** at random among all those of P_n . What does \mathcal{T}_n look like for large n ?

Non-crossing trees

Let \mathcal{T}_n be a random **non-crossing** tree, chosen **uniformly** at random among all those of P_n . What does \mathcal{T}_n look like for large n ?

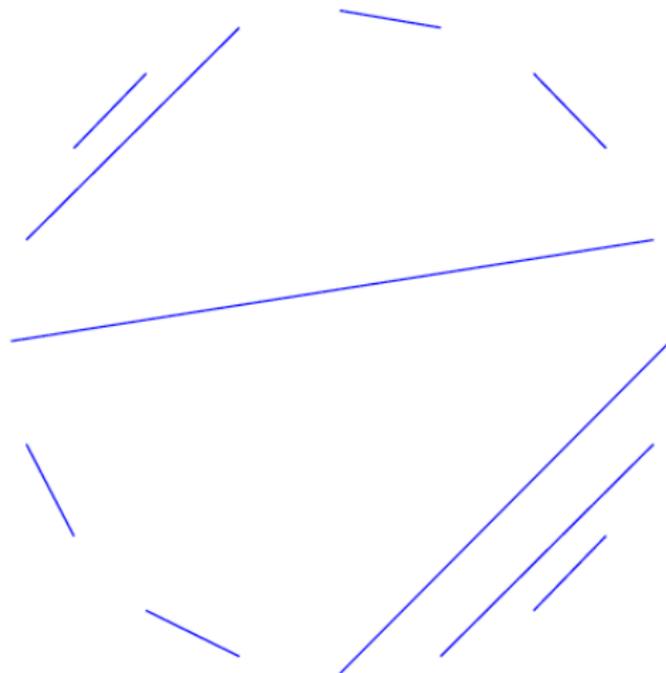


Samples of \mathcal{T}_{500} and \mathcal{T}_{1000} .

Case of **non-crossing** pair-partitions of P_{2n} .

Non-crossing pair partitions

Example of a **non-crossing** pair-partition of P_{20} :

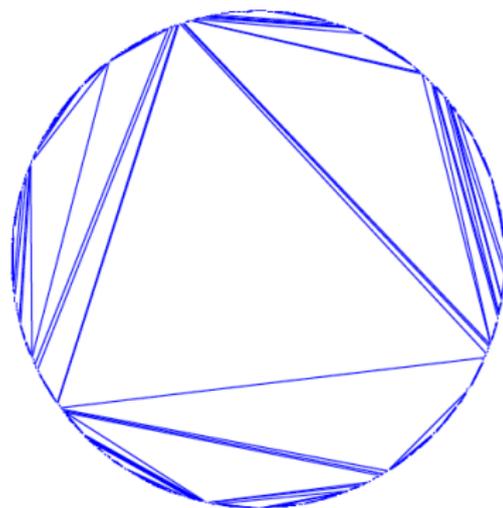
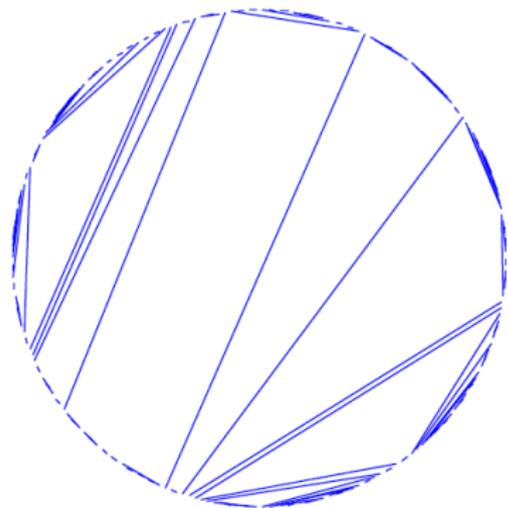


Non-crossing pair partitions

Let \mathcal{Q}_n be a random **non-crossing** pair-partition of P_{2n} , chosen **uniformly** among all those of P_{2n} . What does \mathcal{Q}_n look like for n large ?

Non-crossing pair partitions

Let \mathcal{Q}_n be a random **non-crossing** pair-partition of P_{2n} , chosen **uniformly** among all those of P_{2n} . What does \mathcal{Q}_n look like for n large?



Samples of \mathcal{Q}_{250} and \mathcal{Q}_{1000} .

History of non-crossing configurations of P_n

Combinatorial point of view:

- ▶ Counting and bijections for non-crossing trees: Dulucq & Penaud (1993), Noy (1998), ...
- ▶ Counting of various non-crossing configurations: Flajolet & Noy (1999)

History of non-crossing configurations of P_n

Combinatorial point of view:

- ▶ Counting and bijections for non-crossing trees: Dulucq & Penaud (1993), Noy (1998), ...
- ▶ Counting of various non-crossing configurations: Flajolet & Noy (1999)

Probabilistical combinatorics point of view:

- ▶ Uniform triangulations (maximal degree): Devroye, Flajolet, Hurtado, Noy & Steiger (1999) et Gao & Wormald (2000)
- ▶ Non-crossing trees (total length, maximal degree): Deutsch & Noy (2002), Marckert & Panholzer (2002)
- ▶ Uniform dissections (degrees, maximal degree): Bernasconi, Panagiotou & Steger (2010)

History of non-crossing configurations of P_n

Combinatorial point of view:

- ▶ Counting and bijections for non-crossing trees: Dulucq & Penaud (1993), Noy (1998), ...
- ▶ Counting of various non-crossing configurations: Flajolet & Noy (1999)

Probabilistical combinatorics point of view:

- ▶ Uniform triangulations (maximal degree): Devroye, Flajolet, Hurtado, Noy & Steiger (1999) et Gao & Wormald (2000)
- ▶ Non-crossing trees (total length, maximal degree): Deutsch & Noy (2002), Marckert & Panholzer (2002)
- ▶ Uniform dissections (degrees, maximal degree): Bernasconi, Panagiotou & Steger (2010)

Geometrical point of view:

- ▶ Aldous (1994): large uniform triangulations
- ▶ K' (2011): dissections with large faces (non uniform)

II. CONSTRUCTION OF THE CONTINUOUS LIMITING OBJECT: the **Brownian triangulation** (Aldous, '94)

Interlude: Brownian motion and the Brownian excursion

Brownian motion

Theorem (Donsker)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$.

Brownian motion

Theorem (Donsker)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$.

Brownian motion

Theorem (Donsker)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \right) \xrightarrow[n \rightarrow \infty]{(d)}$$

Brownian motion

Theorem (Donsker)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \right) \xrightarrow[n \rightarrow \infty]{(d)} (W_t, t \geq 0),$$

Brownian motion

Theorem (Donsker)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \right) \xrightarrow[n \rightarrow \infty]{(d)} (W_t, t \geq 0),$$

where $(W_t, t \geq 0)$ is a continuous random function called Brownian motion (which does not depend on σ).

Brownian motion

Theorem (Donsker)

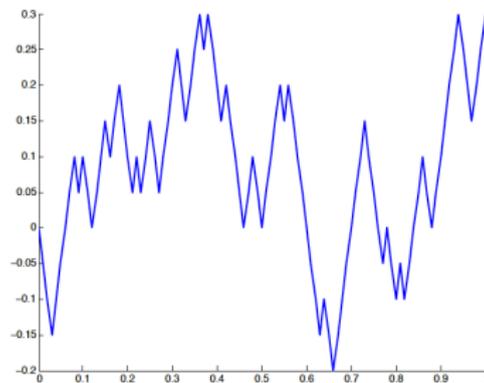
Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \right) \xrightarrow[n \rightarrow \infty]{(d)} (W_t, t \geq 0),$$

where $(W_t, t \geq 0)$ is a continuous random function called Brownian motion (which does not depend on σ).

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, 0 \leq t \leq 1 \right)$$

for $n = 100$:



Brownian motion

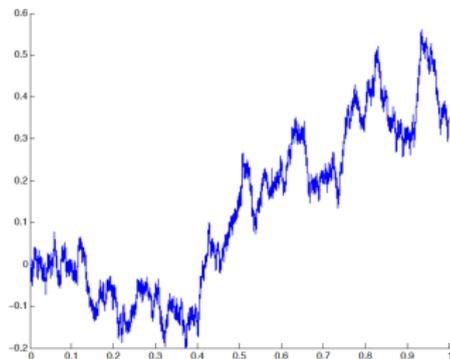
Theorem (Donsker)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \right) \xrightarrow[n \rightarrow \infty]{(d)} (W_t, t \geq 0),$$

where $(W_t, t \geq 0)$ is a continuous random function called *Brownian motion* (which does not depend on σ).

$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, 0 \leq t \leq 1 \right)$
for $n = 100.000$:



Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$.

Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$.

Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \mid S_n = 0, S_i \geq 0 \text{ for } i < n \right) \xrightarrow[n \rightarrow \infty]{(d)}$$

Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \mid S_n = 0, S_i \geq 0 \text{ for } i < n \right) \xrightarrow[n \rightarrow \infty]{(d)} (\mathbb{e}_t, t \geq 0),$$

Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \mid S_n = 0, S_i \geq 0 \text{ for } i < n \right) \xrightarrow[n \rightarrow \infty]{(d)} (e_t, t \geq 0),$$

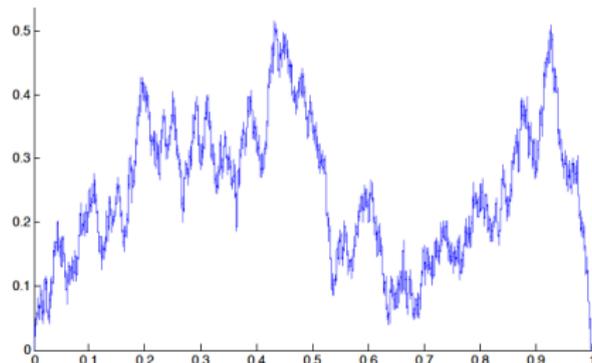
where $(e_t, t \geq 0)$ is a continuous random function called the Brownian excursion.

Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \mid S_n = 0, S_i \geq 0 \text{ for } i < n \right) \xrightarrow[n \rightarrow \infty]{(d)} (e_t, t \geq 0),$$

where $(e_t, t \geq 0)$ is a continuous random function called the Brownian excursion.

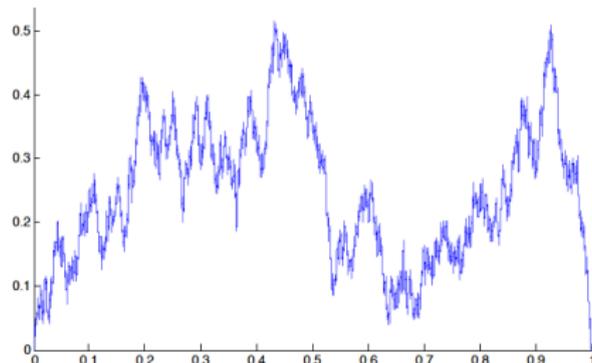


Theorem (Donsker, conditioned version)

Let $(X_n)_{n \geq 1}$ be a sequence of i.i.d. random variables with $\mathbb{E}[X_1] = 0$ and $\sigma^2 = \mathbb{E}[X_1^2] < \infty$. Set $S_n = X_1 + X_2 + \dots + X_n$. Then:

$$\left(\frac{S_{nt}}{\sigma\sqrt{n}}, t \geq 0 \mid S_n = 0, S_i \geq 0 \text{ for } i < n \right) \xrightarrow[n \rightarrow \infty]{(d)} (e_t, t \geq 0),$$

where $(e_t, t \geq 0)$ is a continuous random function called the Brownian excursion.



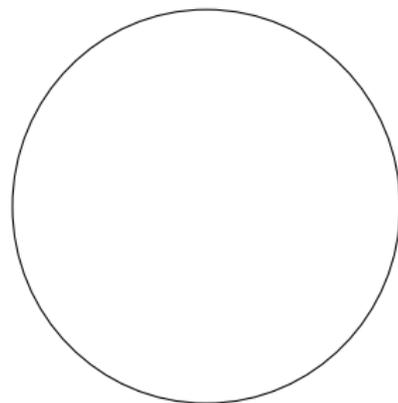
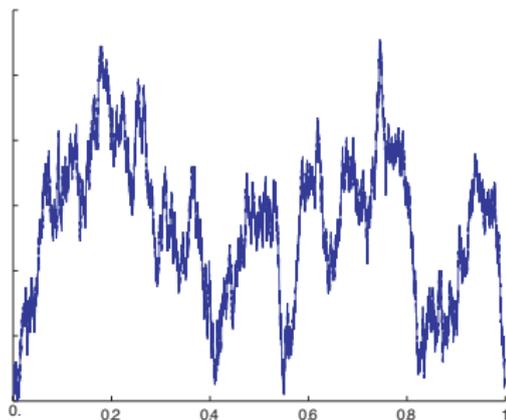
The Brownian excursion can be seen as Brownian motion $(W_t, 0 \leq t \leq 1)$ conditioned on $W_1 = 0$ and $W_t > 0$ for $t \in (0, 1)$.

Construction of the limiting object

We start from the Brownian excursion e :

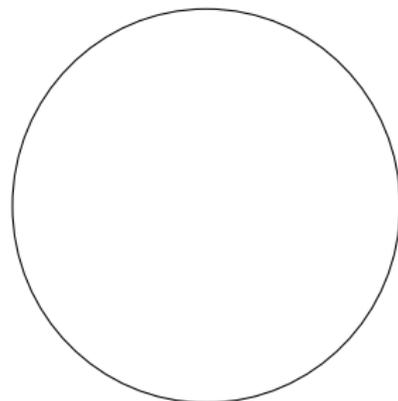
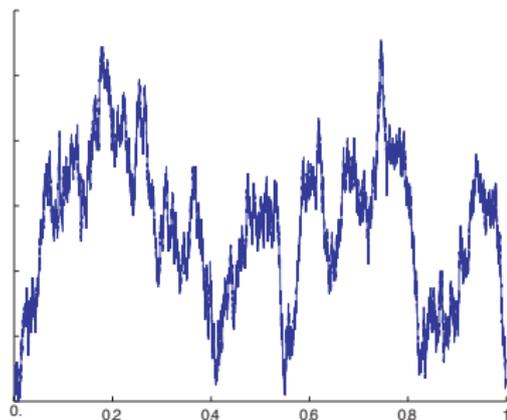
Construction of the limiting object

We start from the Brownian excursion e :



Construction of the limiting object

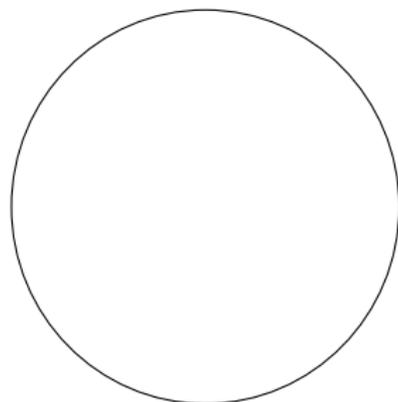
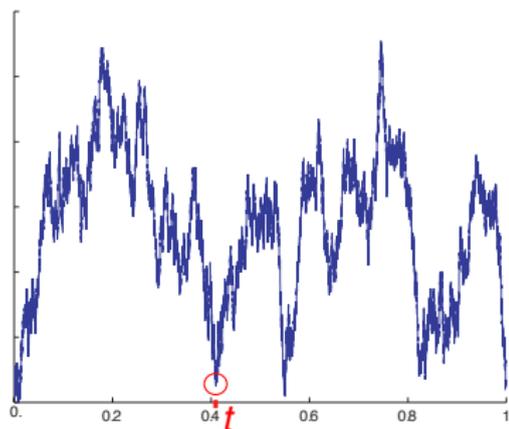
We start from the Brownian excursion e :



Let t be a local minimum time.

Construction of the limiting object

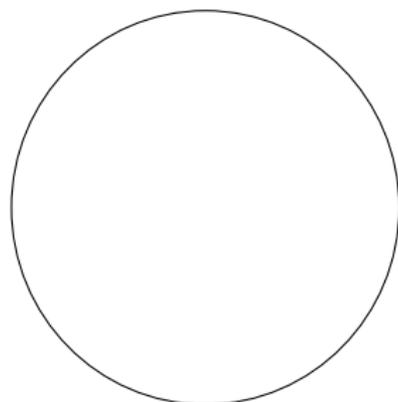
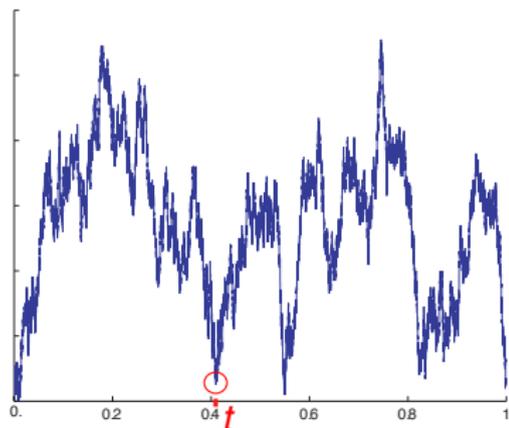
We start from the Brownian excursion e :



Let t be a local minimum time.

Construction of the limiting object

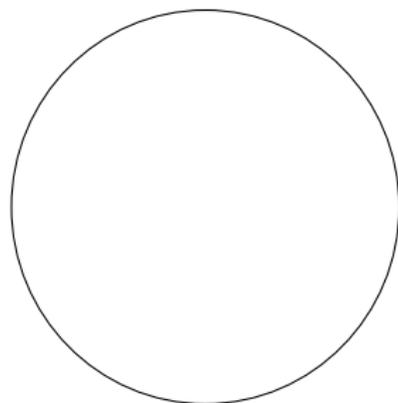
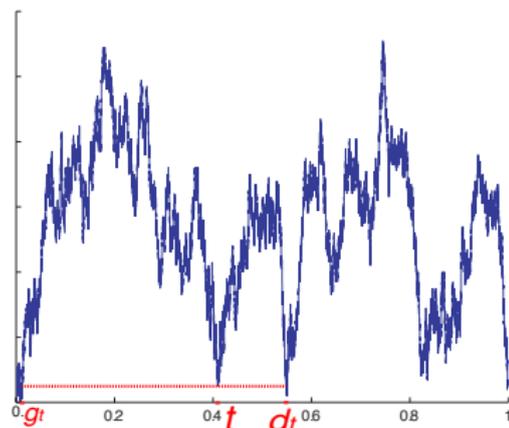
We start from the Brownian excursion e :



Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$.

Construction of the limiting object

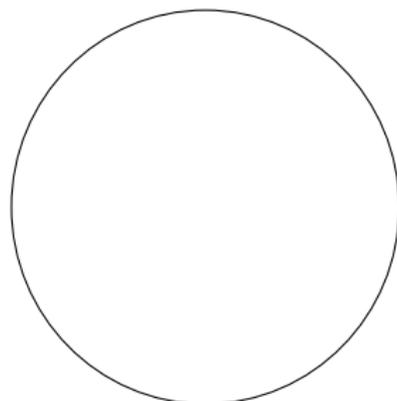
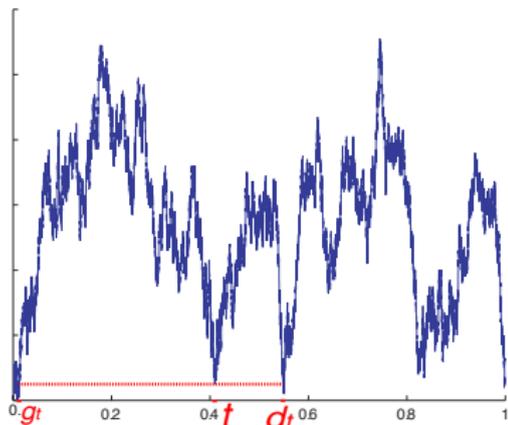
We start from the Brownian excursion e :



Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$.

Construction of the limiting object

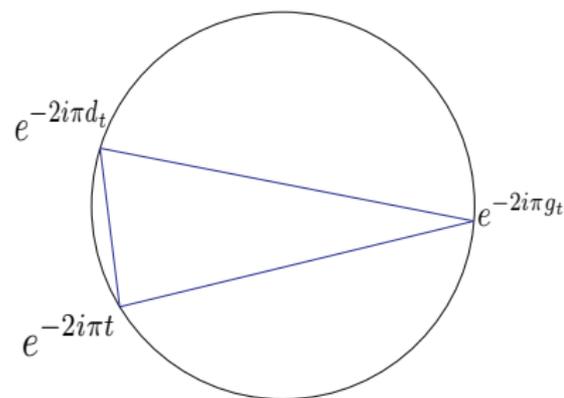
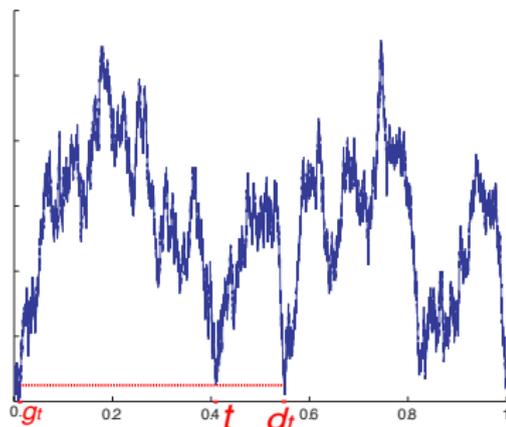
We start from the Brownian excursion e :



Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$. Then draw the chords $[e^{-2i\pi g_t}, e^{-2i\pi t}]$, $[e^{-2i\pi t}, e^{-2i\pi d_t}]$ and $[e^{-2i\pi g_t}, e^{-2i\pi d_t}]$.

Construction of the limiting object

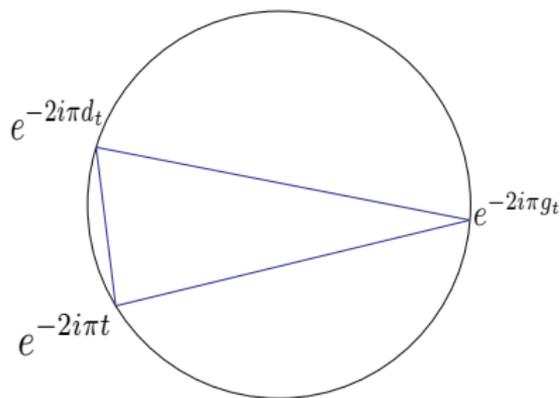
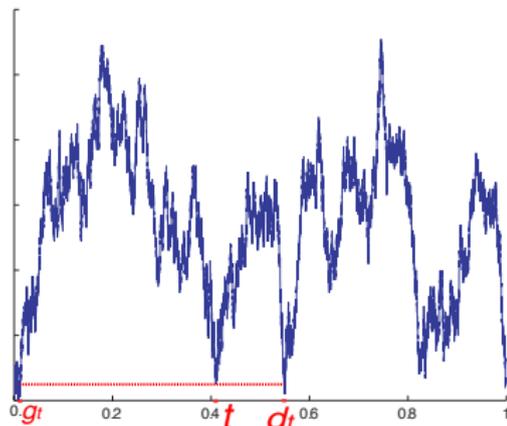
We start from the Brownian excursion e :



Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$. Then draw the chords $[e^{-2i\pi g_t}, e^{-2i\pi t}]$, $[e^{-2i\pi t}, e^{-2i\pi d_t}]$ and $[e^{-2i\pi g_t}, e^{-2i\pi d_t}]$.

Construction of the limiting object

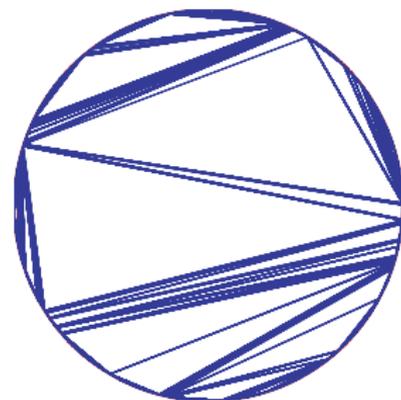
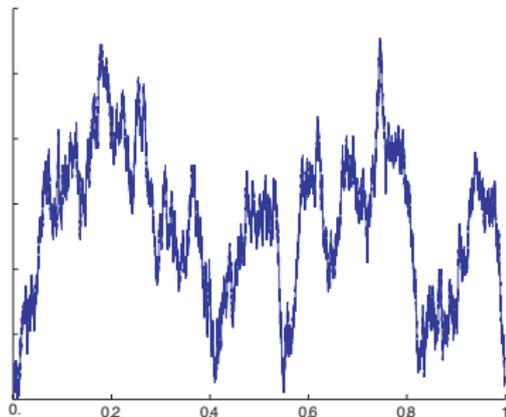
We start from the Brownian excursion e :



Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$. Then draw the chords $[e^{-2i\pi g_t}, e^{-2i\pi t}]$, $[e^{-2i\pi t}, e^{-2i\pi d_t}]$ and $[e^{-2i\pi g_t}, e^{-2i\pi d_t}]$. Repeat this operation for all local minimum times.

Construction of the limiting object

We start from the Brownian excursion e :

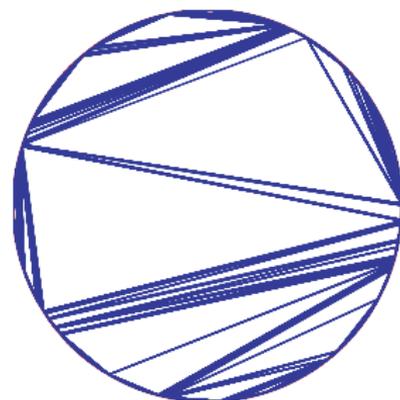
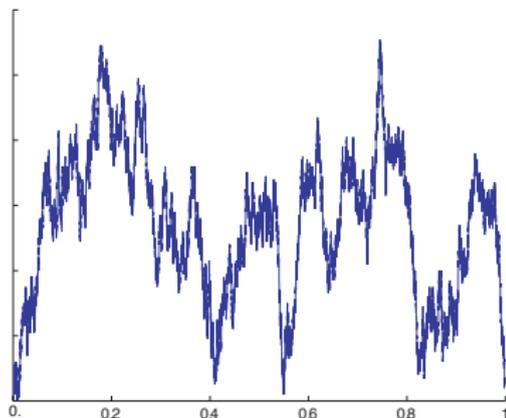


Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$. Then draw the chords $[e^{-2i\pi g_t}, e^{-2i\pi t}]$, $[e^{-2i\pi t}, e^{-2i\pi d_t}]$ and $[e^{-2i\pi g_t}, e^{-2i\pi d_t}]$.

Repeat this operation for all local minimum times.

Construction of the limiting object

We start from the Brownian excursion e :



Let t be a local minimum time. Set $g_t = \sup\{s < t; e_s = e_t\}$ and $d_t = \inf\{s > t; e_s = e_t\}$. Then draw the chords $[e^{-2i\pi g_t}, e^{-2i\pi t}]$, $[e^{-2i\pi t}, e^{-2i\pi d_t}]$ and $[e^{-2i\pi g_t}, e^{-2i\pi d_t}]$.

Repeat this operation for all local minimum times.

The closure of the set thus obtained, denoted by $L(e)$, is called the **Brownian triangulation**.

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} .

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(\mathfrak{e}),$$

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(\mathfrak{e}),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly* distributed dissection of P_n , or a *uniformly* distributed *non-crossing* tree of P_n or a *uniformly* distributed *non-crossing* pair-partition of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(e),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Remarks:

- ▶ Aldous '94: this holds when χ_n is a *uniformly* distributed triangulation of P_n .

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(\mathfrak{e}),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Remarks:

- ▶ Aldous '94: this holds when χ_n is a *uniformly distributed triangulation* of P_n .
- ▶ There exists a “stable” analog of $L(\mathfrak{e})$ with big holes (K. '11).

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(\mathfrak{e}),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Applications:

- ▶ The length of the longest diagonal of χ_n converges in distribution towards the probability measure with density:

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(e),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Applications:

- ▶ The length of the longest diagonal of χ_n converges in distribution towards the probability measure with density:

$$\frac{1}{\pi} \frac{3x-1}{x^2(1-x)^2\sqrt{1-2x}} \mathbf{1}_{\{\frac{1}{3} \leq x \leq \frac{1}{2}\}} dx.$$

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(\mathfrak{e}),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Applications:

- ▶ The length of the longest diagonal of χ_n converges in distribution towards the probability measure with density:

$$\frac{1}{\pi} \frac{3x-1}{x^2(1-x)^2\sqrt{1-2x}} \mathbf{1}_{\{\frac{1}{3} \leq x \leq \frac{1}{2}\}} dx.$$

This stems from a small calculation when χ_n is a triangulation (Aldous '94)!

Theorem (Curien & K. '12)

For $n \geq 3$, let χ_n be a *uniformly distributed dissection* of P_n , or a *uniformly distributed non-crossing tree* of P_n or a *uniformly distributed non-crossing pair-partition* of P_{2n} . Then:

$$\chi_n \xrightarrow[n \rightarrow \infty]{(d)} L(\mathfrak{e}),$$

where the convergence holds in distribution for the Hausdorff distance on compact subsets of the unit disk.

Applications:

- ▶ The length of the longest diagonal of χ_n converges in distribution towards the probability measure with density:

$$\frac{1}{\pi} \frac{3x-1}{x^2(1-x)^2\sqrt{1-2x}} \mathbf{1}_{\{\frac{1}{3} \leq x \leq \frac{1}{2}\}} dx.$$

This stems from a small calculation when χ_n is a triangulation (Aldous '94)!

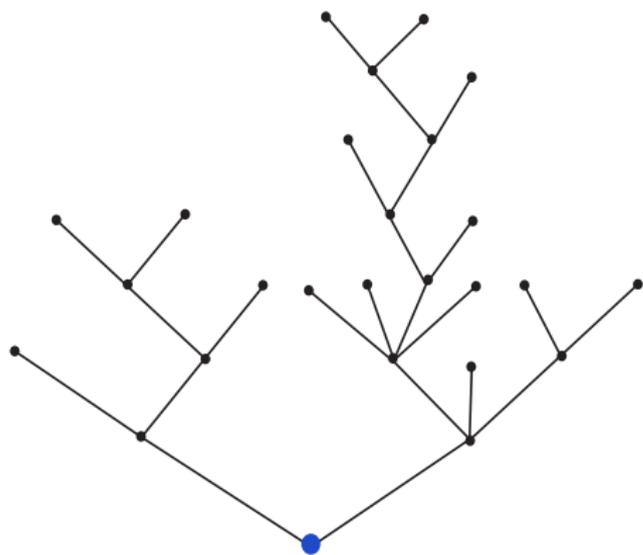
- ▶ The area of the largest face of χ_n converges in distribution towards the area of the largest triangle of $L(\mathfrak{e})$.

**III. HOW DOES ONE ESTABLISH THE CONVERGENCE OF ALL THESE
NON-CROSSING UNIFORMLY DISTRIBUTED MODELS TOWARDS THE
BROWNIAN TRIANGULATION?**

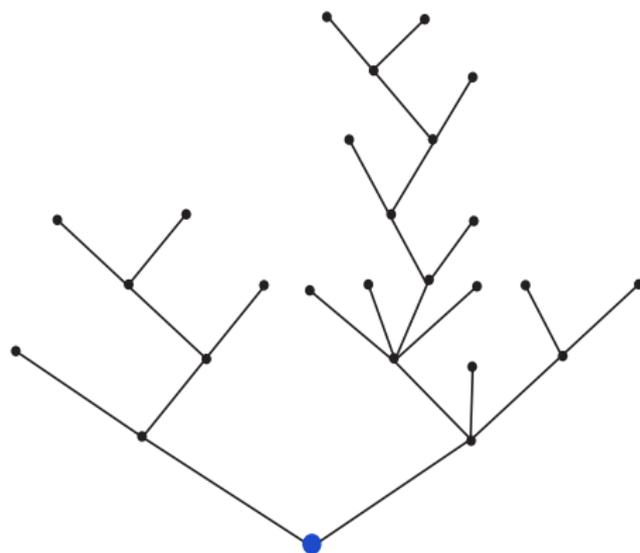
III. HOW DOES ONE ESTABLISH THE CONVERGENCE OF ALL THESE NON-CROSSING UNIFORMLY DISTRIBUTED MODELS TOWARDS THE BROWNIAN TRIANGULATION?

Key point: Each one of the previous models can be coded by a **conditioned Galton-Watson tree**.

Coding trees



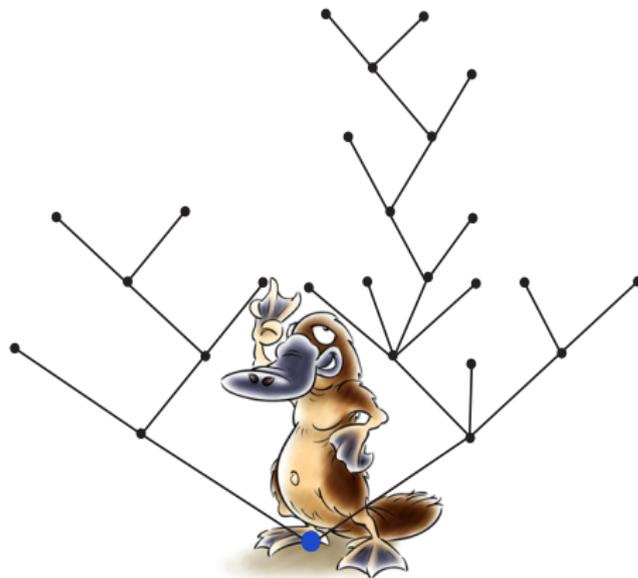
Coding trees



Definition (of the contour function)

A platypus explores the tree at unit speed. For $0 \leq t \leq 2(\zeta(\tau) - 1)$, $C_t(\tau)$ is defined as the distance from the root at the position of the beast at time t .

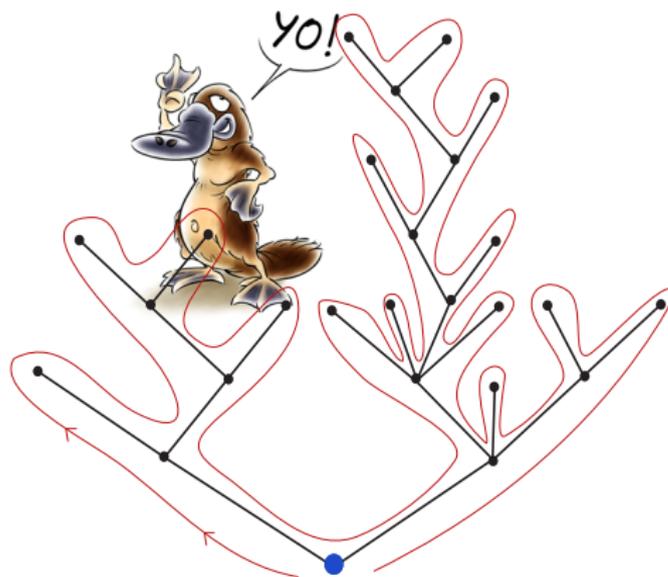
Coding trees



Definition (of the contour function)

A platypus explores the tree at unit speed. For $0 \leq t \leq 2(\zeta(\tau) - 1)$, $C_t(\tau)$ is defined as the distance from the root at the position of the beast at time t .

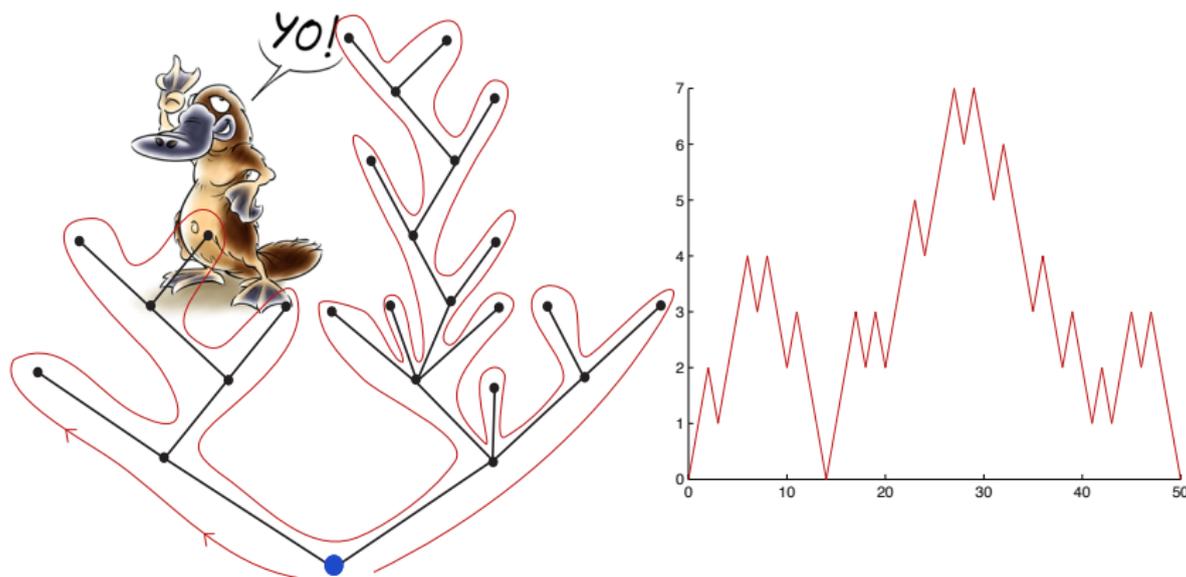
Coding trees



Definition (of the contour function)

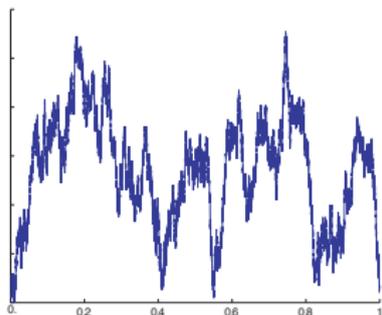
A platypus explores the tree at unit speed. For $0 \leq t \leq 2(\zeta(\tau) - 1)$, $C_t(\tau)$ is defined as the distance from the root at the position of the beast at time t .

Coding trees

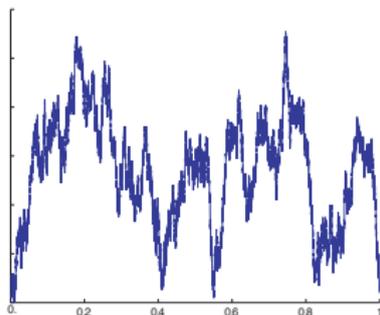


Definition (of the contour function)

A platypus explores the tree at unit speed. For $0 \leq t \leq 2(\zeta(\tau) - 1)$, $C_t(\tau)$ is defined as the distance from the root at the position of the beast at time t .

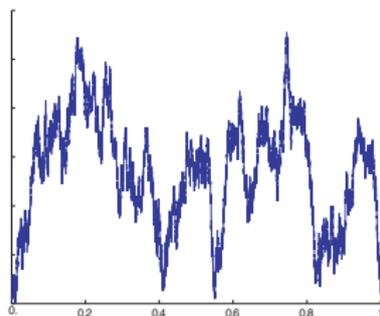


Scaled contour function of a large conditioned Galton-Watson tree.



Scaled contour function of a large conditioned Galton-Watson tree.

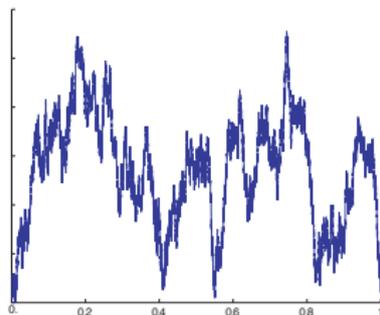
Strategy to prove the convergence towards the Brownian triangulation:



Scaled contour function of a large conditioned Galton-Watson tree.

Strategy to prove the convergence towards the Brownian triangulation:

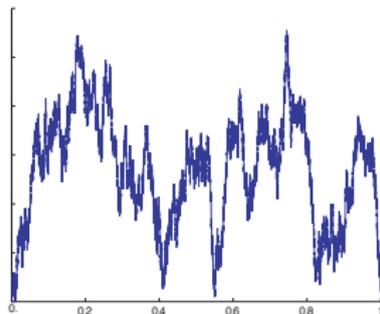
- ▶ Each one of the **non-crossing uniformly distributed** models can be coded by a **conditioned Galton-Watson** tree.



Scaled contour function of a large conditioned Galton-Watson tree.

Strategy to prove the convergence towards the Brownian triangulation:

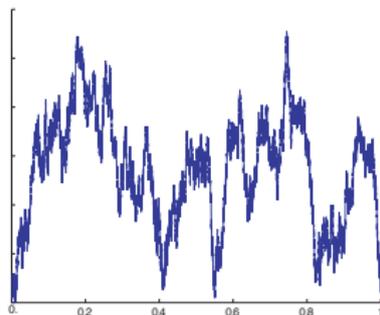
- ▶ Each one of the **non-crossing uniformly distributed** models can be coded by a **conditioned Galton-Watson** tree.
- ▶ The scaled contour functions of **conditioned Galton-Watson** trees converge towards the **Brownian excursion**.



Scaled contour function of a large conditioned Galton-Watson tree.

Strategy to prove the convergence towards the Brownian triangulation:

- ▶ Each one of the **non-crossing uniformly distributed** models can be coded by a **conditioned Galton-Watson** tree.
- ▶ The scaled contour functions of **conditioned Galton-Watson** trees converge towards the **Brownian excursion**.
- ▶ **The Brownian excursion** codes the Brownian triangulation $L(\mathfrak{e})$.



Scaled contour function of a large conditioned Galton-Watson tree.

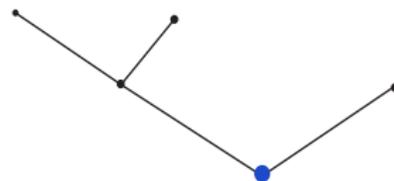
Strategy to prove the convergence towards the Brownian triangulation:

- ▶ Each one of the **non-crossing uniformly distributed** models can be coded by a **conditioned Galton-Watson** tree.
- ▶ The scaled contour functions of **conditioned Galton-Watson** trees converge towards the **Brownian excursion**.
- ▶ **The Brownian excursion** codes the Brownian triangulation $L(e)$.

It follows that the **non-crossing uniformly distributed** models converge towards $L(e)$.

Brief recap on Galton-Watson trees

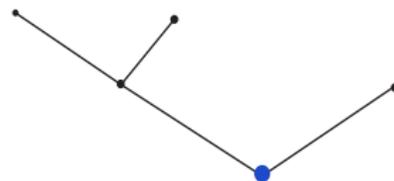
We consider rooted plane (oriented) trees.



Brief recap on Galton-Watson trees

We consider rooted plane (oriented) trees.

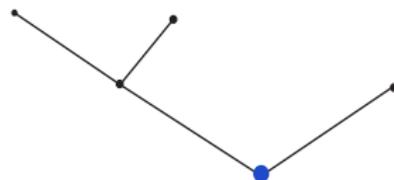
Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$.



Brief recap on Galton-Watson trees

We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

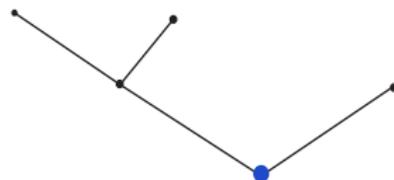


Brief recap on Galton-Watson trees

We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

1. k_\emptyset is distributed according to ρ , where k_\emptyset is the number of children of **the root**.

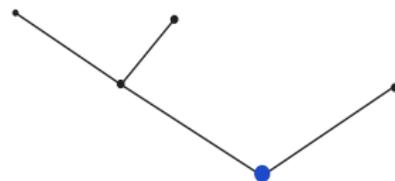


Brief recap on Galton-Watson trees

We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

1. k_\emptyset is distributed according to ρ , where k_\emptyset is the number of children of **the root**.
2. for every $j \geq 1$ with $\rho(j) > 0$, conditionally on $\mathbb{P}_\rho(\cdot | k_\emptyset = j)$, the j subtrees of the j children of **the root** are independent with law \mathbb{P}_ρ .



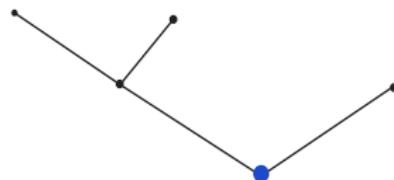
Brief recap on Galton-Watson trees

We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

1. k_\emptyset is distributed according to ρ , where k_\emptyset is the number of children of **the root**.
2. for every $j \geq 1$ with $\rho(j) > 0$, conditionally on $\mathbb{P}_\rho(\cdot | k_\emptyset = j)$, the j subtrees of the j children of **the root** are independent with law \mathbb{P}_ρ .

Here, $k_\emptyset = 2$.



Brief recap on Galton-Watson trees

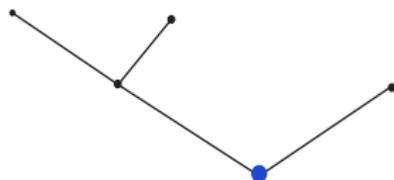
We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

1. k_\emptyset is distributed according to ρ , where k_\emptyset is the number of children of **the root**.
2. for every $j \geq 1$ with $\rho(j) > 0$, conditionally on $\mathbb{P}_\rho(\cdot | k_\emptyset = j)$, the j subtrees of the j children of **the root** are independent with law \mathbb{P}_ρ .

Here, $k_\emptyset = 2$.

The probability of getting this tree is $\rho(2)^2 \rho(0)^3$.



Brief recap on Galton-Watson trees

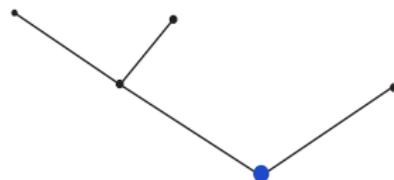
We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

1. k_\emptyset is distributed according to ρ , where k_\emptyset is the number of children of **the root**.
2. for every $j \geq 1$ with $\rho(j) > 0$, conditionally on $\mathbb{P}_\rho(\cdot | k_\emptyset = j)$, the j subtrees of the j children of **the root** are independent with law \mathbb{P}_ρ .

Here, $k_\emptyset = 2$.

The probability of getting this tree is $\rho(2)^2 \rho(0)^3$.



$\zeta(\tau)$ is the total number of vertices and $\lambda(\tau)$ is the total number of leaves.

Brief recap on Galton-Watson trees

We consider rooted plane (oriented) trees.

Let ρ be a probability measure on \mathbb{N} with mean ≤ 1 s.t. $\rho(1) < 1$. The law of a **Galton-Watson** tree with offspring distribution ρ is the unique probability distribution \mathbb{P}_ρ on the set of all trees such that:

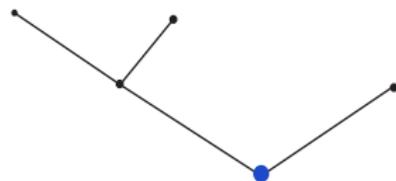
1. k_\emptyset is distributed according to ρ , where k_\emptyset is the number of children of **the root**.
2. for every $j \geq 1$ with $\rho(j) > 0$, conditionally on $\mathbb{P}_\rho(\cdot | k_\emptyset = j)$, the j subtrees of the j children of **the root** are independent with law \mathbb{P}_ρ .

Here, $k_\emptyset = 2$.

The probability of getting this tree is $\rho(2)^2 \rho(0)^3$.

Here, $\zeta(\tau) = 5$ and $\lambda(\tau) = 3$.

$\zeta(\tau)$ is the total number of vertices and $\lambda(\tau)$ is the total number of leaves.



Brief recap on Galton-Watson trees

Proposition

Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW** $_{\nu}$ tree conditioned on having n vertices.

Brief recap on Galton-Watson trees

Proposition

Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW** $_{\nu}$ tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n .

Brief recap on Galton-Watson trees

Proposition

Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW** $_{\nu}$ tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n . We have (k_u being the number of children of u):

$$\mathbb{P}_{\nu}[\tau] = \prod_{u \in \tau} \nu_{k_u}$$

Brief recap on Galton-Watson trees

Proposition

Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW** $_{\nu}$ tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n . We have (k_u being the number of children of u):

$$\mathbb{P}_{\nu}[\tau] = \prod_{u \in \tau} \nu_{k_u} = \prod_{u \in \tau} \frac{1}{2^{k_u+1}}$$

Brief recap on Galton-Watson trees

Proposition

Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW** $_{\nu}$ tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n . We have (k_u being the number of children of u):

$$\mathbb{P}_{\nu}[\tau] = \prod_{u \in \tau} \nu_{k_u} = \prod_{u \in \tau} \frac{1}{2^{k_u+1}} = 2^{-\sum_{u \in \tau} (k_u + 1)}$$

Brief recap on Galton-Watson trees

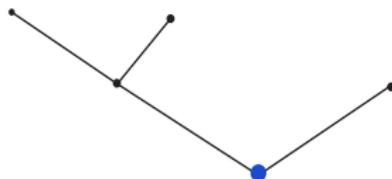
Proposition

Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW $_{\nu}$** tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n . We have (k_u being the number of children of u):

$$\mathbb{P}_{\nu}[\tau] = \prod_{u \in \tau} \nu_{k_u} = \prod_{u \in \tau} \frac{1}{2^{k_u+1}} = 2^{-\sum_{u \in \tau} (k_u + 1)}$$



Brief recap on Galton-Watson trees

Proposition

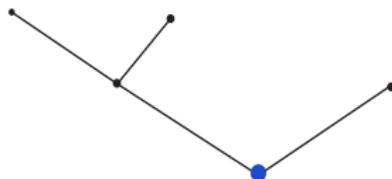
Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW $_{\nu}$** tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n . We have (k_u being the number of children of u):

$$\mathbb{P}_{\nu}[\tau] = \prod_{u \in \tau} \nu_{k_u} = \prod_{u \in \tau} \frac{1}{2^{k_u+1}} = 2^{-\sum_{u \in \tau} (k_u + 1)}$$

$$\begin{aligned} \sum_{u \in \tau} (k_u + 1) &= 3 + 3 + 1 + 1 + 1 = 9 \\ &= 2 \times 5 - 1 \end{aligned}$$



Brief recap on Galton-Watson trees

Proposition

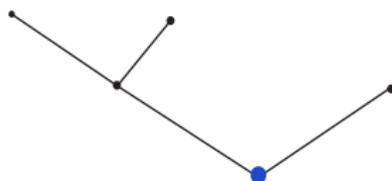
Let ν be defined by $\nu(k) = 1/2^{k+1}$ for $k \geq 0$. Then the law of a **uniformly** distributed tree with n vertices is the law of a **GW $_{\nu}$** tree conditioned on having n vertices.

Proof.

Let τ be a tree with n vertices. It suffices to prove that $\mathbb{P}_{\nu}[\tau]$ depends only on n . We have (k_u being the number of children of u):

$$\mathbb{P}_{\nu}[\tau] = \prod_{u \in \tau} \nu_{k_u} = \prod_{u \in \tau} \frac{1}{2^{k_u+1}} = 2^{-\sum_{u \in \tau} (k_u + 1)} = 2^{-2n+1}.$$

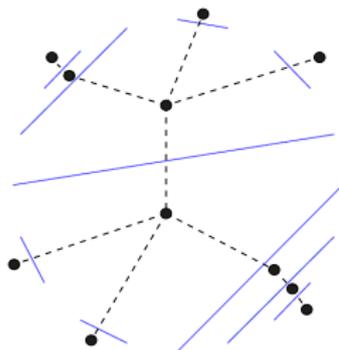
$$\begin{aligned} \sum_{u \in \tau} (k_u + 1) &= 3 + 3 + 1 + 1 + 1 = 9 \\ &= 2 \times 5 - 1 \end{aligned}$$



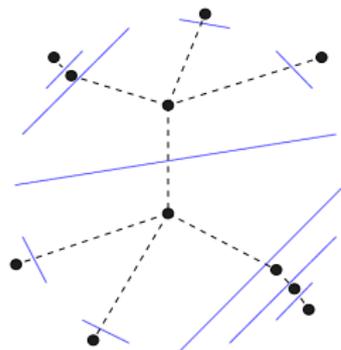
How can one code **non-crossing** **uniformly distributed** models by a **conditioned**
Galton-Watson tree?

Coding **uniform** pair-partitions by **Galton-Watson** trees.

Consider the dual of a **uniform non-crossing** pair-partition of P_{2n} :

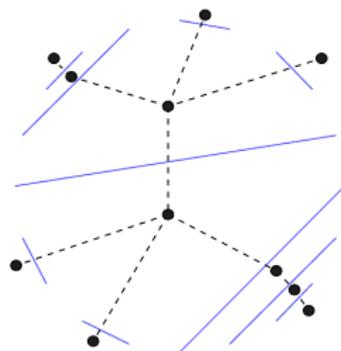


Consider the dual of a **uniform non-crossing** pair-partition of P_{2n} :



It is a **uniform** tree with n edges.

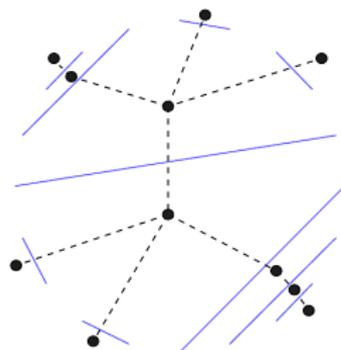
Consider the dual of a **uniform non-crossing** pair-partition of P_{2n} :



It is a **uniform** tree with n edges.

Hence the law of a conditioned **Galton-Watson** tree with offspring distribution $\text{Geom}(1/2)$, conditioned on having n edges.

Consider the dual of a **uniform non-crossing** pair-partition of P_{2n} :



It is a **uniform** tree with n edges.

Hence the law of a conditioned **Galton-Watson** tree with offspring distribution $\text{Geom}(1/2)$, conditioned on having n edges.

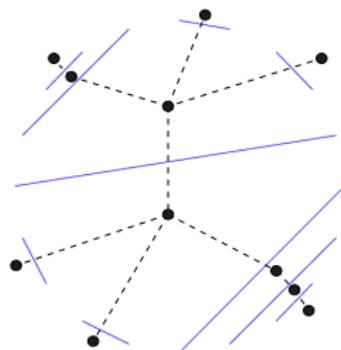
Theorem (Aldous '93)

Let \mathfrak{t}_n be a random tree distributed according to $\mathbb{P}_{\text{Geom}(1/2)}[\cdot \mid \zeta(\tau) = n + 1]$.

Let σ^2 be the variance of $\text{Geom}(1/2)$. Then:

$$\left(\frac{\sigma}{2\sqrt{n}} C_{2nt}(\mathfrak{t}_n), 0 \leq t \leq 1 \right) \xrightarrow[n \rightarrow \infty]{(d)} (\mathbb{e}_t, 0 \leq t \leq 1).$$

Consider the dual of a **uniform non-crossing** pair-partition of P_{2n} :



It is a **uniform** tree with n edges.

Hence the law of a conditioned **Galton-Watson** tree with offspring distribution $\text{Geom}(1/2)$, conditioned on having n edges.

Theorem (Aldous '93)

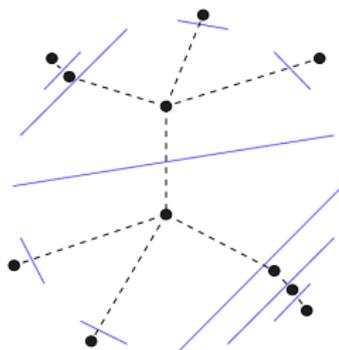
Let \mathfrak{t}_n be a random tree distributed according to $\mathbb{P}_{\text{Geom}(1/2)}[\cdot \mid \zeta(\tau) = n + 1]$.

Let σ^2 be the variance of $\text{Geom}(1/2)$. Then:

$$\left(\frac{\sigma}{2\sqrt{n}} C_{2nt}(\mathfrak{t}_n), 0 \leq t \leq 1 \right) \xrightarrow[n \rightarrow \infty]{(d)} (\mathfrak{e}_t, 0 \leq t \leq 1).$$

Idea: the contour function of a **Galton-Watson** tree behaves as a random walk.

Consider the dual of a **uniform non-crossing** pair-partition of P_{2n} :



It is a **uniform** tree with n edges.

Hence the law of a conditioned **Galton-Watson** tree with offspring distribution $\text{Geom}(1/2)$, conditioned on having n edges.

Theorem (Aldous '93)

Let \mathfrak{t}_n be a random tree distributed according to $\mathbb{P}_{\text{Geom}(1/2)}[\cdot \mid \zeta(\tau) = n + 1]$.

Let σ^2 be the variance of $\text{Geom}(1/2)$. Then:

$$\left(\frac{\sigma}{2\sqrt{n}} C_{2nt}(\mathfrak{t}_n), 0 \leq t \leq 1 \right) \xrightarrow[n \rightarrow \infty]{(d)} (\mathfrak{e}_t, 0 \leq t \leq 1).$$

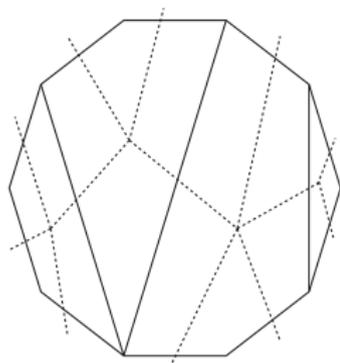
Idea: the contour function of a **Galton-Watson** tree behaves as a random walk.

It follows that **uniform non-crossing** pair-partitions of P_{2n} converge towards **the Brownian triangulation**.

Coding **uniform** dissections by **Galton-Watson** trees.

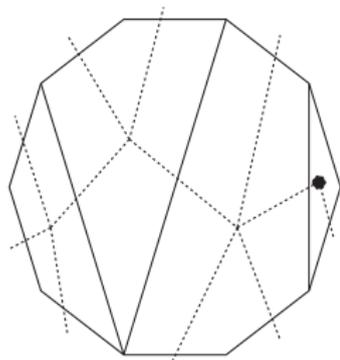
Coding uniform dissections by Galton-Watson trees

Consider the dual of a **uniform** dissection of P_n :



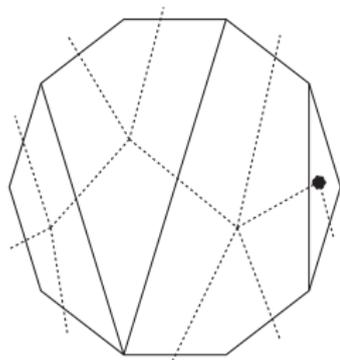
Coding uniform dissections by Galton-Watson trees

Consider the dual of a **uniform** dissection of P_n , suitably rooted:



Coding uniform dissections by Galton-Watson trees

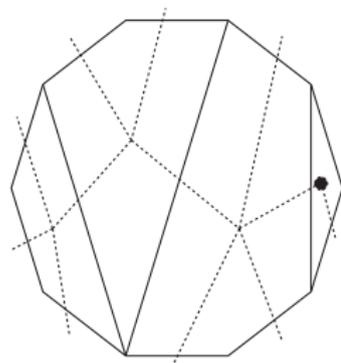
Consider the dual of a **uniform** dissection of P_n , suitably rooted:



This is a **uniform** tree on the set of all trees with $n - 1$ leaves s.t. no vertex has exactly one child.

Coding uniform dissections by Galton-Watson trees

Consider the dual of a **uniform** dissection of P_n , suitably rooted:



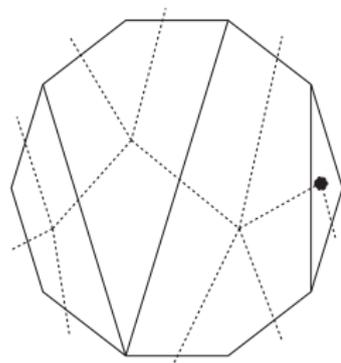
This is a **uniform** tree on the set of all trees with $n - 1$ leaves s.t. no vertex has exactly one child.

Proposition (Curien & K. '12, Pitman & Rizzolo '11)

The law of a **uniform** tree on the set of all trees with $n - 1$ leaves s.t. no vertex has exactly one child is the law of a GW_{μ_0} tree with offspring distribution μ_0 conditioned on having $n - 1$ leaves, where:

Coding uniform dissections by Galton-Watson trees

Consider the dual of a **uniform** dissection of P_n , suitably rooted:



This is a **uniform** tree on the set of all trees with $n - 1$ leaves s.t. no vertex has exactly one child.

Proposition (Curien & K. '12, Pitman & Rizzolo '11)

The law of a **uniform** tree on the set of all trees with $n - 1$ leaves s.t. no vertex has exactly one child is the law of a GW_{μ_0} tree with offspring distribution μ_0 conditioned on having $n - 1$ leaves, where:

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1} \text{ for } i \geq 2.$$

Coding uniform dissections by Galton-Watson trees

Theorem (K. '11)

Let \mathfrak{t}_n be a random tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n]$. Let σ^2 be the variance of μ_0 . Then:

$$\left(\frac{\sigma}{2\sqrt{\zeta(\mathfrak{t}_n)}} C_{2\zeta(\mathfrak{t}_n)t}(\mathfrak{t}_n), 0 \leq t \leq 1 \right) \xrightarrow[n \rightarrow \infty]{(d)} (\mathfrak{e}_t, 0 \leq t \leq 1).$$

Coding uniform dissections by Galton-Watson trees

Theorem (K. '11)

Let \mathbf{t}_n be a random tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n]$. Let σ^2 be the variance of μ_0 . Then:

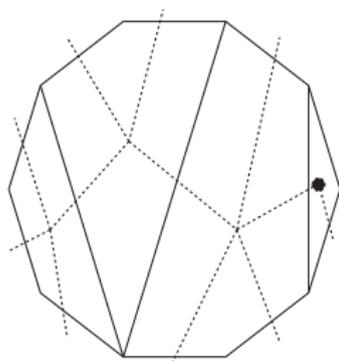
$$\left(\frac{\sigma}{2\sqrt{\zeta(\mathbf{t}_n)}} C_{2\zeta(\mathbf{t}_n)t}(\mathbf{t}_n), 0 \leq t \leq 1 \right) \xrightarrow[n \rightarrow \infty]{(d)} (\mathbf{e}_t, 0 \leq t \leq 1).$$

It follows that **uniform dissections** of P_n converge towards **the Brownian triangulation**.

Conclusion: In these **uniform** models, some **independence** is hiding.

IV. APPLICATION TO THE STUDY OF **UNIFORM** DISSECTIONS

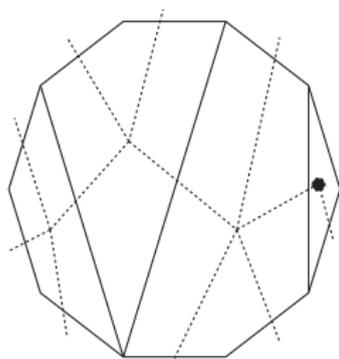
Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
 the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
 where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
 the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
 where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

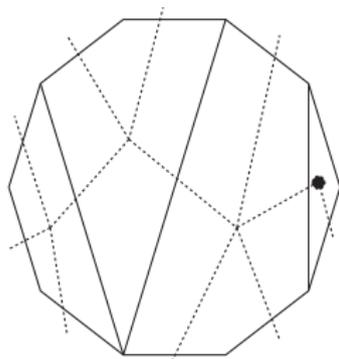
Application 1 (Counting dissections). Probabilistic proof of the following result:

Theorem (Flajolet & Noy '99)

Let a_n be the number of dissections of P_n . Then:

$$a_n \underset{n \rightarrow \infty}{\sim}$$

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

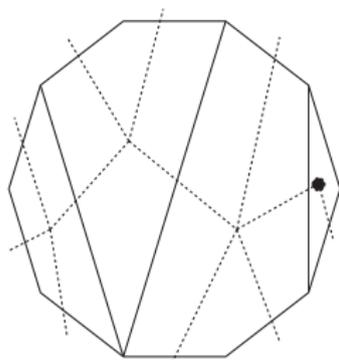
Application 1 (Counting dissections). Probabilistic proof of the following result:

Theorem (Flajolet & Noy '99)

Let a_n be the number of dissections of P_n . Then:

$$a_n \underset{n \rightarrow \infty}{\sim} \frac{1}{4} \sqrt{\frac{99\sqrt{2} - 140}{\pi}} n^{-3/2} (3 + 2\sqrt{2})^n.$$

Application to the study of uniform dissections

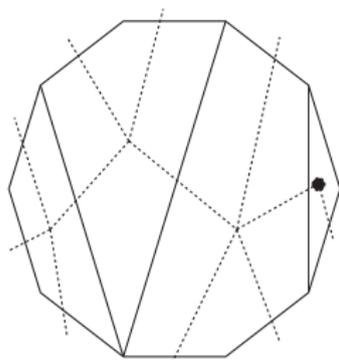


\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
 the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
 where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

Application 2 (Study of the maximal face degree). Denote by $D^{(n)}$ the maximal face degree of \mathcal{D}_n .

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0} [\cdot | \lambda(\tau) = n - 1]$,
where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

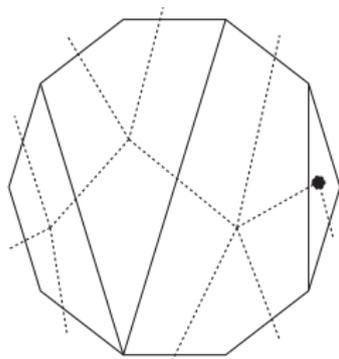
Application 2 (Study of the maximal face degree). Denote by $D^{(n)}$ the maximal face degree of \mathcal{D}_n .

Theorem (Curien & K. '12)

Set $\beta = 2 + \sqrt{2}$. For every $c > 0$, we have:

$$\mathbb{P}(\log_{\beta}(n) - c \log_{\beta} \log_{\beta}(n) \leq D^{(n)} \leq \log_{\beta}(n) + c \log_{\beta} \log_{\beta}(n)) \xrightarrow[n \rightarrow \infty]{} 1.$$

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
 the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
 where ($i \geq 2$):

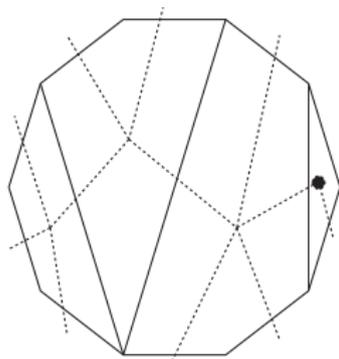
$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

Application 3 (Study of the vertex degree).

Theorem (Curien & K. '12)

Let $\partial^{(n)}$ be the number of diagonals ending at the vertex with affix 1 in \mathcal{D}_n .

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

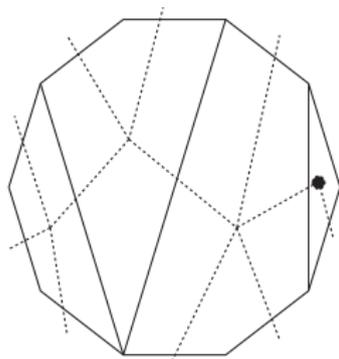
Application 3 (Study of the vertex degree).

Theorem (Curien & K. '12)

Let $\partial^{(n)}$ be the number of diagonals ending at the vertex with affix 1 in \mathcal{D}_n .
Then $\partial^{(n)}$ converges in distribution towards the sum of two independent
 $\text{Geom}(\sqrt{2} - 1)$ random variables, i.e. for $k \geq 0$:

$$\mathbb{P}(\partial^{(n)} = k) \xrightarrow[n \rightarrow \infty]{} (k + 1)\mu_0^2(1 - \mu_0)^k.$$

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
 the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
 where ($i \geq 2$):

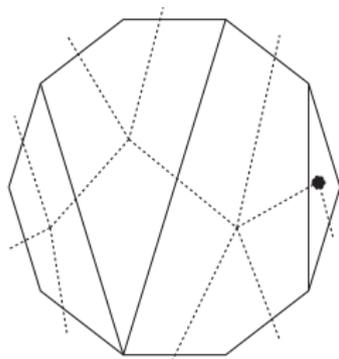
$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

Application 4 (Study of the maximal vertex degree). Proof of a conjecture
 by Bernasconi, Panagiotou & Steger:

Theorem (Curien & K. '12)

Let $\Delta^{(n)}$ be the maximal number of diagonals ending at any vertex in \mathcal{D}_n .

Application to the study of uniform dissections



\mathcal{D}_n : **uniform** dissection of P_n . Recall that:
the dual of \mathcal{D}_n is a tree with law $\mathbb{P}_{\mu_0}[\cdot | \lambda(\tau) = n - 1]$,
where ($i \geq 2$):

$$\mu_0(0) = \frac{2 - \sqrt{2}}{2}, \quad \mu_0(1) = 0, \quad \mu_0(i) = (2 - \sqrt{2})^{i-1}.$$

Application 4 (Study of the maximal vertex degree). Proof of a conjecture by Bernasconi, Panagiotou & Steger:

Theorem (Curien & K. '12)

Let $\Delta^{(n)}$ be the maximal number of diagonals ending at any vertex in \mathcal{D}_n . Set $b = \sqrt{2} + 1$. Then for every $c > 0$, we have

$$\mathbb{P}(\Delta^{(n)} \geq \log_b(n) + (1 + c) \log_b \log_b(n)) \xrightarrow[n \rightarrow \infty]{} 0.$$

Conjecture

Let $\Delta^{(n)}$ be the maximum number of diagonals ending at some vertex of \mathcal{D}_n .
Set $b = \sqrt{2} + 1$. For every $c > 0$:

$$\mathbb{P} \left(\left| \Delta^{(n)} - (\log_b(n) + \log_b \log_b(n)) \right| > c \log_b \log_b(n) \right) \xrightarrow[n \rightarrow \infty]{} 0.$$

This is satisfied for another value of b in the case of uniform triangulations
(Devroye, Flajolet, Hurtado, Noy & Steiger '99 et Gao & Wormald '00)

Conjecture

Let $\Delta^{(n)}$ be the maximum number of diagonals ending at some vertex of \mathcal{D}_n .
Set $b = \sqrt{2} + 1$. For every $c > 0$:

$$\mathbb{P} \left(\left| \Delta^{(n)} - (\log_b(n) + \log_b \log_b(n)) \right| > c \log_b \log_b(n) \right) \xrightarrow[n \rightarrow \infty]{} 0.$$

This is satisfied for another value of b in the case of uniform triangulations
(Devroye, Flajolet, Hurtado, Noy & Steiger '99 et Gao & Wormald '00)

