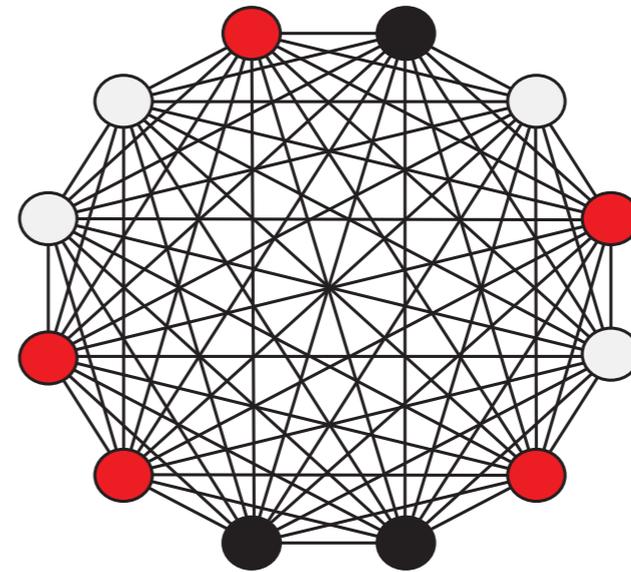
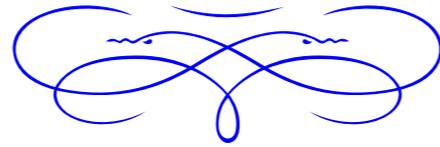


On a *prey-predator* model



Igor Kortchemski
CNRS & CMAP, École polytechnique

Bonn probability seminar – July 2015

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,
- at fixed rate 1, each **predator** propagates to every neighbouring **prey**.

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,
- at fixed rate 1, each **predator** propagates to every neighbouring **prey**.

Motivations :

 Model of two competing species, or model of first-passage percolation with destruction.

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,
- at fixed rate 1, each **predator** propagates to every neighbouring **prey**.

Motivations :

 Model of two competing species, or model of first-passage percolation with destruction.

 Other possible analogies:

vacant vertex \longleftrightarrow

prey \longleftrightarrow

predator \longleftrightarrow

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,
- at fixed rate 1, each **predator** propagates to every neighbouring **prey**.

Motivations :

 Model of two competing species, or model of first-passage percolation with destruction.

 Other possible analogies:

vacant vertex	\longleftrightarrow	vacant vertex
prey	\longleftrightarrow	healthy cell
predator	\longleftrightarrow	cell infected by a virus

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,
- at fixed rate 1, each **predator** propagates to every neighbouring **prey**.

Motivations :

 Model of two competing species, or model of first-passage percolation with destruction.

 Other possible analogies:

vacant vertex	\longleftrightarrow	normal individual
prey	\longleftrightarrow	individual trying to spread a rumor (spreader)
predator	\longleftrightarrow	individual trying to scotch the rumor (stifler)

What is this about?

On a graph, we are interested in the following **prey-predator** model (introduced by Bordenave '12) :

- each vertex is either occupied by a **prey**, or a **predator**, or is **vacant**,
- at fixed rate $\lambda > 0$, each **prey** propagates to every **vacant** neighbour,
- at fixed rate 1, each **predator** propagates to every neighbouring **prey**.

Motivations :

 Model of two competing species, or model of first-passage percolation with destruction.

 Other possible analogies:

vacant vertex	\longleftrightarrow	Susceptible (S) individual
prey	\longleftrightarrow	Infected (I) individual
predator	\longleftrightarrow	Recovered (R) individual

Here, $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$.

Other type of models studied in the literature:

Here, $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$.

Other type of models studied in the literature:

- SIR model (Kermack—McKendrick '27), where $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $I \xrightarrow{1} R$

Here, $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$.

Other type of models studied in the literature:

- SIR model (Kermack—McKendrick '27), where $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $I \xrightarrow{1} R$
- Daley—Kendall ('65) rumour propagation model, where $\{I, S\} \xrightarrow{1} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$, $\{I, I\} \xrightarrow{1} \{R, R\}$.

Here, $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$.

Other type of models studied in the literature:

- SIR model (Kermack—McKendrick '27), where $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $I \xrightarrow{1} R$
- Daley–Kendall ('65) rumour propagation model, where
 $\{I, S\} \xrightarrow{1} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$, $\{I, I\} \xrightarrow{1} \{R, R\}$.
- Maki–Thompson ('73) directed rumour propagation model, where
 $(I, S) \xrightarrow{1} (I, I)$, $(R, I) \xrightarrow{1} (R, R)$, $(I, I) \xrightarrow{1} (I, R)$.

Here, $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$.

Other type of models studied in the literature:

- SIR model (Kermack—McKendrick '27), where $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $I \xrightarrow{1} R$
- Daley–Kendall ('65) rumour propagation model, where
 $\{I, S\} \xrightarrow{1} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$, $\{I, I\} \xrightarrow{1} \{R, R\}$.
- Maki–Thompson ('73) directed rumour propagation model, where
 $(I, S) \xrightarrow{1} (I, I)$, $(R, I) \xrightarrow{1} (R, R)$, $(I, I) \xrightarrow{1} (I, R)$.
- Williams Bjercknes ('71) tumor growth model (or biased voter model),
 where $(I, S) \xrightarrow{\lambda} (I, I)$, $(S, I) \xrightarrow{1} (S, S)$.

Here, $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$.

Other type of models studied in the literature:

- SIR model (Kermack—McKendrick '27), where $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $I \xrightarrow{1} R$
- Daley–Kendall ('65) rumour propagation model, where $\{I, S\} \xrightarrow{1} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$, $\{I, I\} \xrightarrow{1} \{R, R\}$.
- Maki–Thompson ('73) directed rumour propagation model, where $(I, S) \xrightarrow{1} (I, I)$, $(R, I) \xrightarrow{1} (R, R)$, $(I, I) \xrightarrow{1} (I, R)$.
- Williams Bjercknes ('71) tumor growth model (or biased voter model), where $(I, S) \xrightarrow{\lambda} (I, I)$, $(S, I) \xrightarrow{1} (S, S)$.
- Kordzakhia ('05), where $\{I, S\} \xrightarrow{\lambda} \{I, I\}$, $\{R, I\} \xrightarrow{1} \{R, R\}$, $\{R, S\} \xrightarrow{1} \{R, R\}$.

FACEBOOK

Facebook Is About to Lose 80% of Its Users, Study Says

Social media is like a disease that spreads, and then dies

By Sam Frizell @Sam_Frizell | Jan. 21, 2014 | 266 Comments

 Share

 Tweet 4,155

 Share 741

 *Pin it*

 Read Later

Facebook's growth will eventually come to a quick end, much like an infectious disease that spreads rapidly and suddenly dies, say Princeton researchers who are using diseases to model the life cycles of social media.



Facebook To Lose 80% Of Users By 2017

InformationWeek - 23 janv. 2014

Online social networks spread like disease epidemics, and **Facebook** will **lose** 80% of its victims -- I mean, users -- by 2017, according to a study from Princeton University researchers. The study, "Epidemiological modeling of online social network dynamics" ...

Facebook could lose 80 percent of users by 2017, report claims

Fox News - 23 janv. 2014

"**Facebook** has already reached the peak of its popularity and has entered a decline phase," they concluded. "The future suggests that **Facebook** will undergo a rapid decline in the coming years, **losing** 80 percent of its peak user base between 2015 and 2017 ...

Facebook will lose 80 percent of its users in next 4 years, Princeton study says

The Star-Ledger - NJ.com - 23 janv. 2014

Most of the 874 million people across the world who sign on to **Facebook** will stop doing so in the next four years, according to a Princeton University study. The study predicts the social media site will **lose** 80 percent of the users it had at its 2010 peak ...

Facebook Losing Users; 30 Years of Mac Ads; Snapchat 'Ghost' Verification

PC Magazine - 23 janv. 2014

Topping tech headlines Wednesday, a new study predicts a rapid decline for **Facebook**, which researchers said will **lose** 80 percent of its peak user base between 2015 and 2017. Using epidemiological models to track the spread of infectious diseases and ...



Facebook Might Lose 80% of Users and be the Next 'MySpace,' Study Says

Morning Ledger - 23 janv. 2014

Facebook Might **Lose** 80 percent and be the Next MySpace A new study conducted and released by Princeton University has described social networks as similar to infectious diseases. It pointed out that such sites gain millions of users within just a short span ...



Facebook Will Lose 80 Percent of Users by 2017

Guardian Liberty Voice - 23 janv. 2014

Facebook According to researchers at Princeton University, **Facebook** will **lose** 80 percent of its users by 2017. The researchers have also stated that that decline is already happening now and could reach the total any time within 2015 and the 2017 deadline.



Facebook to 'lose 80% of users by 2017'

Irish Times - 23 janv. 2014

Facebook has spread like an infectious disease but we are slowly becoming immune to its attractions, and the platform will be largely abandoned by 2017, say researchers at Princeton University. The forecast of **Facebook's** impending doom was made by ...

Facebook will LOSE 80% of its users by 2017 – epidemiological study

Register - 23 janv. 2014

According to the students' paper, **Facebook** is "just beginning to show the onset of an abandonment phase", after reaching its popularity peak in 2012, which will lead to it **losing** 80 per cent of its peak user base between 2015 and 2017. The paper, which has ...

Facebook Predicts Princeton Won't Exist In 2021

InformationWeek - 24 janv. 2014

Princeton's report, from the university's Department of Mechanical and Aerospace Engineering, used Google search data to predict engagement trends, ultimately concluding that **Facebook** was set to **lose** a whopping 80% of users by 2017. Such a ...

Could Facebook Really Lose 80% of its Users?

DailyFinance - 23 janv. 2014

Facebook has so far been the only super-hot social media network to escape the fate of former top sites like MySpace, Friendster, or even GeoCities/Tripod back in the day. And with its now-successful stock offering and seeming ubiquity among nearly every ...



Facebook Will Lose 80 Percent Of Users In Next Three Years, Researchers Say

Opposing Views - 23 janv. 2014

People are slowly building up an immunity to **Facebook** and researchers predict it will **lose** 80 percent of its peak user base by 2017. Researchers at Princeton University compared the growth of the social media site to the spread of disease. They believe ...

Outline

I. TEST YOUR INTUITION!

Outline

I. TEST YOUR INTUITION!

II. PREYS & PREDATORS ON A COMPLETE GRAPH

Outline

- I. TEST YOUR INTUITION!**
- II. PREYS & PREDATORS ON A COMPLETE GRAPH**
- III. PREYS & PREDATORS ON AN INFINITE TREE**

I. TEST YOUR INTUITION!



II. PREYS & PREDATORS ON A COMPLETE GRAPH

III. PREYS & PREDATORS ON AN INFINITE TREE

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls.

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls. You take out each time a ball at random.

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball.

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball. If the ball was blue, you put it back in the box and take out a red ball.

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball. If the ball was blue, you put it back in the box and take out a red ball.

You keep doing it until left only with balls of the same color. How many balls will be left (as a function of n)?

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball. If the ball was blue, you put it back in the box and take out a red ball.

You keep doing it until left only with balls of the same color. How many balls will be left (as a function of n)?

- 1) Roughly ϵn for some $\epsilon > 0$.
- 2) Roughly \sqrt{n} .
- 3) Roughly $\log n$.
- 4) Roughly a constant.
- 5) Some other behavior.

Seen on Gil Kalai's blog

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball. If the ball was blue, you put it back in the box and take out a red ball.

You keep doing it until left only with balls of the same color. How many balls will be left (as a function of n)?

- 1) Roughly ϵn for some $\epsilon > 0$.
- 2) Roughly \sqrt{n} .
- 3) Roughly $\log n$.
- 4) Roughly a constant.
- 5) Some other behavior.

Other formulation (O.K. Corral problem, Williams & McIlroy, 1998) . There are two groups of n gunmen that shoot at each other. Once a gunman is hit he stops shooting, and leaves the place happily and peacefully. How many gunmen will be left after all gunmen in one team have left?



Figure: Excerpt of the film "Gunfight at the O.K. Corral" (1957)

Vu sur le blog de Gil Kalai

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball. If the ball was blue, you put it back in the box and take out a red ball.

You keep as before until left only with balls of the same color. How many balls will be left (as a function of n)?

- 1) Roughly ϵn for some $\epsilon > 0$.
- 2) Roughly \sqrt{n} .
- 3) Roughly $\log n$.
- 4) Roughly a constant.
- 5) Some other behavior.

Other formulation (O.K. Corral problem, Williams & McIlroy, 1998) . There are two groups of n gunmen that shoot at each other. Once a gunman is hit he stops shooting, and leaves the place happily and peacefully. How many gunmen will be left after all gunmen in one team have left?

Vu sur le blog de Gil Kalai

You have a box with n red balls and n blue balls. You take out each time a ball at random. If the ball was red, you put it back in the box and take out a blue ball. If the ball was blue, you put it back in the box and take out a red ball.

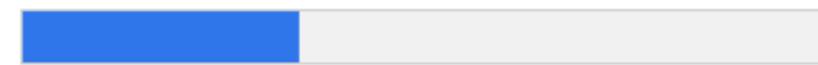
You keep as before until left only with balls of the same color. How many balls will be left (as a function of n)?

- 1) Roughly ϵn for some $\epsilon > 0$.
- 2) Roughly \sqrt{n} .
- 3) Roughly $\log n$.
- 4) Roughly a constant.
- 5) Some other behavior.

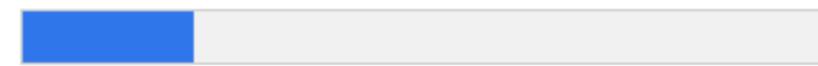
Other formulation (O.K. Corral problem): two groups of n gunmen that shoot at each other until they stop shooting, and leaves the place having only one group of k gunmen. How many k will be left after all gunmen in one team

How many balls will be left when you take out a ball of the opposite color

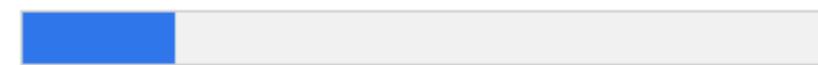
A constant time n 34.62% (45 votes)



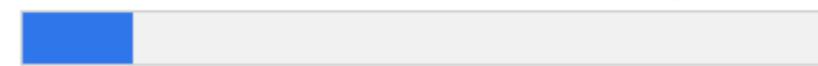
square root n 21.54% (28 votes)



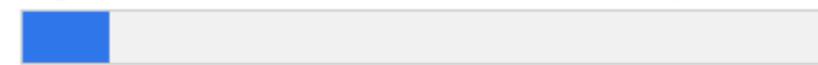
A constant 19.23% (25 votes)



Some other behavior 13.85% (18 votes)



$\log n$ 11% (14 votes)



Total Votes: 130

There are k hit he
y gunmen

Kingman & Volkov's solution (1/3)

If urn **A** has m balls and urn **B** has n balls, the probability that a ball is removed from **A** is $\frac{n}{m+n}$.

Kingman & Volkov's solution (1/3)

If urn **A** has m balls and urn **B** has n balls, the probability that a ball is removed from **A** is $\frac{n}{m+n}$. But

$$\frac{n}{m+n} = \frac{1/m}{1/m + 1/n} = \mathbb{P}(\text{Exp}(1/m) < \text{Exp}(1/n)).$$

Kingman & Volkov's solution (2/3)

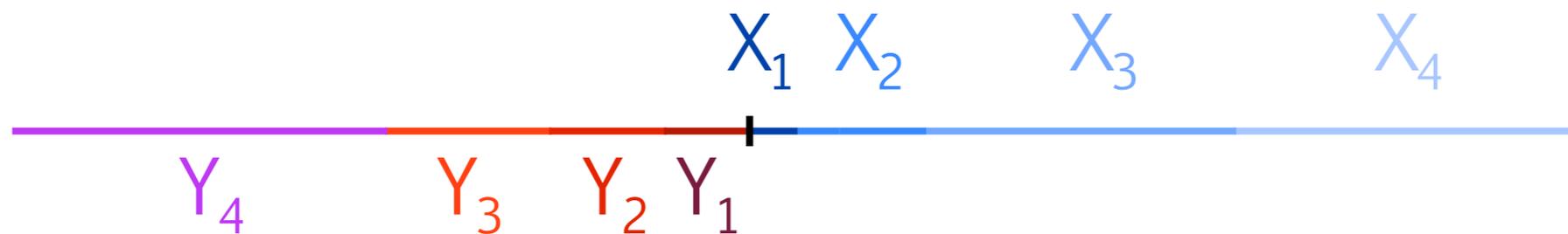
Let $(X_i, Y_i)_{i \geq 1}$ be independent random variables such that X_i are Y_i exponential random variables with **mean** i .

Kingman & Volkov's solution (2/3)

Let $(X_i, Y_i)_{i \geq 1}$ be independent random variables such that X_i are Y_i exponential random variables with **mean** i .

Consider a piece of wood represented by the interval $[-n, n]$ and made of $2n$ pieces such that

$$\text{length}([i-1, i]) = X_i, \quad \text{length}([-i, -i+1]) = Y_i \quad (1 \leq i \leq n).$$

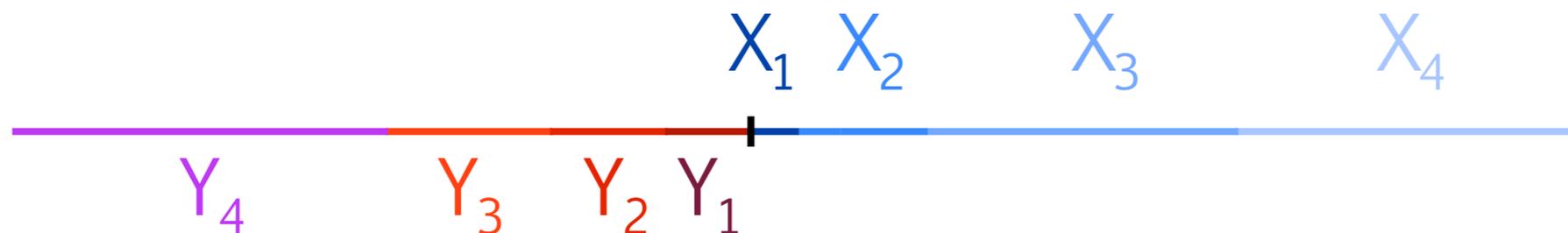


Kingman & Volkov's solution (2/3)

Let $(X_i, Y_i)_{i \geq 1}$ be independent random variables such that X_i are Y_i exponential random variables with **mean** i .

Consider a piece of wood represented by the interval $[-n, n]$ and made of $2n$ pieces such that

$$\text{length}([i-1, i]) = X_i, \quad \text{length}([-i, -i+1]) = Y_i \quad (1 \leq i \leq n).$$



Light both ends, and stop the fire when the origin is reached.

Kingman & Volkov's solution (2/3)

Let $(X_i, Y_i)_{i \geq 1}$ be independent random variables such that X_i are Y_i exponential random variables with **mean** i .

Consider a piece of wood represented by the interval $[-n, n]$ and made of $2n$ pieces such that

$$\text{length}([i-1, i]) = X_i, \quad \text{length}([-i, -i+1]) = Y_i \quad (1 \leq i \leq n).$$



Light both ends, and stop the fire when the origin is reached. Let $R(n)$ be the number of remaining pieces.

Kingman & Volkov's solution (2/3)

Let $(X_i, Y_i)_{i \geq 1}$ be independent random variables such that X_i are Y_i exponential random variables with **mean** i .

Consider a piece of wood represented by the interval $[-n, n]$ and made of $2n$ pieces such that

$$\text{length}([i-1, i]) = X_i, \quad \text{length}([-i, -i+1]) = Y_i \quad (1 \leq i \leq n).$$



Light both ends, and stop the fire when the origin is reached. Let $R(n)$ be the number of remaining pieces. Then $R(n)$ has the same law as the number of remaining balls in the urn/gunman problem.

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Set $S_k = X_1 + \cdots + X_k$.

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Set $S_k = X_1 + \cdots + X_k$. We have $\mathbb{E}[S_k] \simeq k^2$

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Set $S_k = X_1 + \cdots + X_k$. We have $\mathbb{E}[S_k] \simeq k^2$, so $S_k \simeq k^2$.

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Set $S_k = X_1 + \dots + X_k$. We have $\mathbb{E}[S_k] \simeq k^2$, so $S_k \simeq k^2$. But, if the **left** part burns first, $S_{R(n)} \simeq L(n)$.

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Set $S_k = X_1 + \dots + X_k$. We have $\mathbb{E}[S_k] \simeq k^2$, so $S_k \simeq k^2$. But, if the **left** part burns first, $S_{R(n)} \simeq L(n)$. Hence

$$R(n)^2 \simeq n^{3/2}$$

Kingman & Volkov's solution (3/3)

In order to estimate the number $R(n)$ of remaining pieces, first estimate the remaining length $L(n)$:

$$L(n) = \left| \sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right|.$$

Then

$$\text{Var} \left(\sum_{i=1}^n X_i - \sum_{i=1}^n Y_i \right) = \sum_{j=1}^n 2j^2 \simeq n^3.$$

Hence

$$L(n) \simeq n^{3/2}.$$

Set $S_k = X_1 + \dots + X_k$. We have $\mathbb{E}[S_k] \simeq k^2$, so $S_k \simeq k^2$. But, if the **left** part burns first, $S_{R(n)} \simeq L(n)$. Hence

$$R(n)^2 \simeq n^{3/2}$$

so that $R(n) \simeq n^{3/4}$.

This “decoupling” idea is called the Athreya–Karlin embedding, and is useful to study more general Pólya urn schemes.

I. TEST YOUR INTUITION!

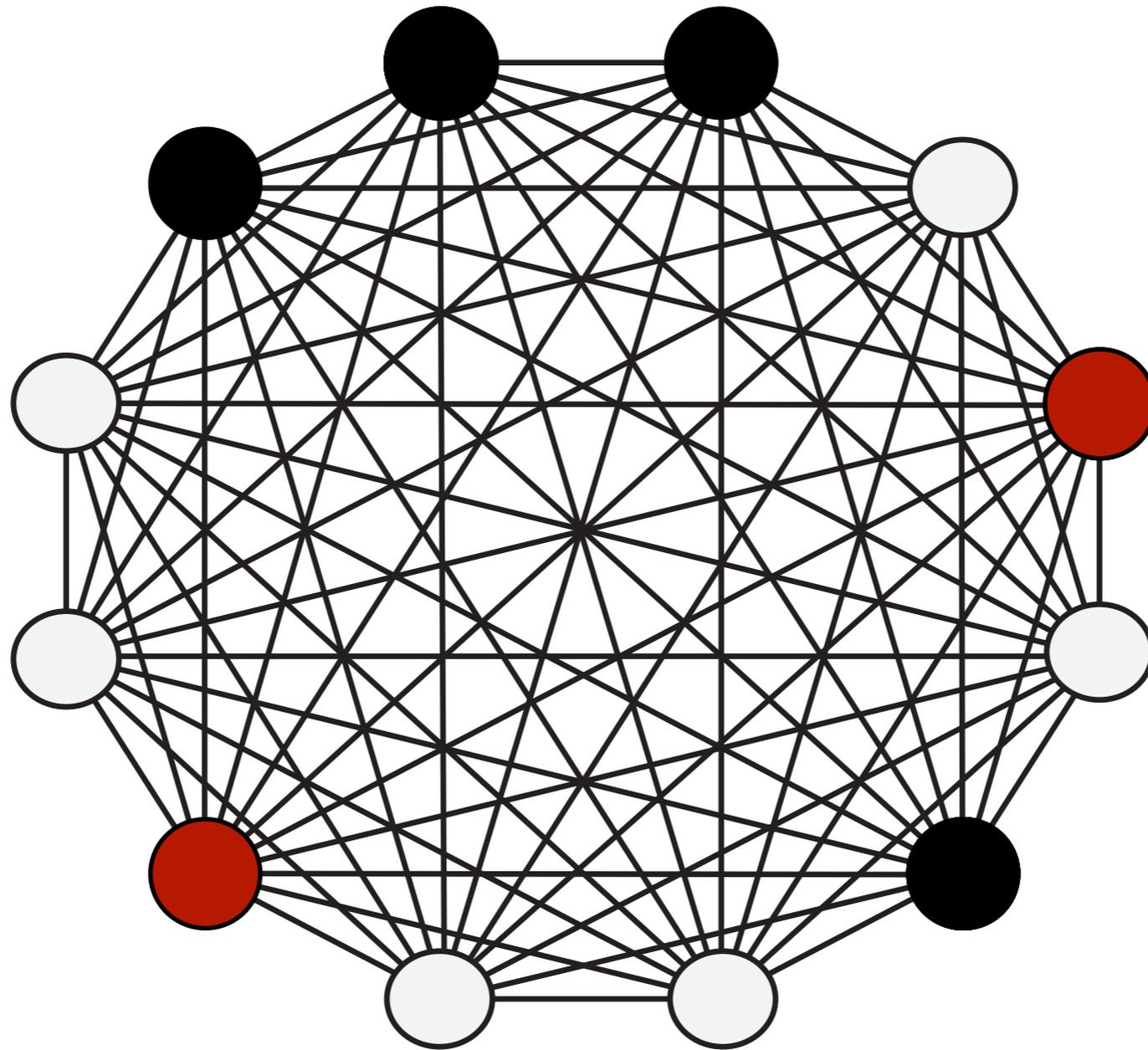
II. PREY & PREDATORS ON A COMPLETE GRAPH



III. PREYS & PREDATORS ON AN INFINITE TREE

We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.

We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.



We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.

Set

$$E_{\text{ext}}^N = \{\text{at a certain moment, there are no more } \mathbf{S} \text{ vertices}\}.$$

We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.

Set

$$E_{\text{ext}}^N = \{\text{at a certain moment, there are no more } \mathbf{S} \text{ vertices}\}.$$

Question. How does $\mathbb{P}(E_{\text{ext}}^N)$ behave as $N \rightarrow \infty$?

We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.

Set

$$E_{\text{ext}}^N = \{\text{at a certain moment, there are no more } S \text{ vertices}\}.$$

Question. How does $\mathbb{P}(E_{\text{ext}}^N)$ behave as $N \rightarrow \infty$?

Theorem (K. '13).

We have

$$\mathbb{P}(E_{\text{ext}}^N) \xrightarrow{N \rightarrow \infty} \begin{cases} 0 & \text{if } \lambda \in \\ & \text{if } \lambda = \\ 1 & \text{if } \lambda > . \end{cases}$$

We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.

Set

$$E_{\text{ext}}^N = \{\text{at a certain moment, there are no more } S \text{ vertices}\}.$$

Question. How does $\mathbb{P}(E_{\text{ext}}^N)$ behave as $N \rightarrow \infty$?

Theorem (K. '13).

We have

$$\mathbb{P}(E_{\text{ext}}^N) \xrightarrow{N \rightarrow \infty} \begin{cases} 0 & \text{if } \lambda \in (0, 1) \\ & \text{if } \lambda = 1 \\ 1 & \text{if } \lambda > 1. \end{cases}$$

We consider K_{N+2} , a complete graph on $N + 2$ vertices, and start the dynamics with one **I** vertex, one **R** vertex and N **S** vertices.

Set

$$E_{\text{ext}}^N = \{\text{at a certain moment, there are no more } S \text{ vertices}\}.$$

Question. How does $\mathbb{P}(E_{\text{ext}}^N)$ behave as $N \rightarrow \infty$?

Theorem (K. '13).

We have

$$\mathbb{P}(E_{\text{ext}}^N) \xrightarrow{N \rightarrow \infty} \begin{cases} 0 & \text{if } \lambda \in (0, 1) \\ \frac{1}{2} & \text{if } \lambda = 1 \\ 1 & \text{if } \lambda > 1. \end{cases}$$

DECOUPLING USING YULE PROCESSES



Transition rates

Let S_t, I_t, R_t be the population sizes at time t .

Total rate of $\{S, I\} \rightarrow \{I, I\}$:

Transition rates

Let S_t, I_t, R_t be the population sizes at time t .

$$\text{Total rate of } \{S, I\} \rightarrow \{I, I\} \quad : \quad \lambda \cdot S_t \cdot I_t.$$

Transition rates

Let S_t, I_t, R_t be the population sizes at time t .

Total rate of $\{S, I\} \rightarrow \{I, I\}$: $\lambda \cdot S_t \cdot I_t$.

Total rate $\{R, I\} \rightarrow \{R, R\}$:

Transition rates

Let S_t, I_t, R_t be the population sizes at time t .

$$\text{Total rate of } \{S, I\} \rightarrow \{I, I\} \quad : \quad \lambda \cdot S_t \cdot I_t.$$

$$\text{Total rate } \{R, I\} \rightarrow \{R, R\} \quad : \quad I_t \cdot R_t.$$

Transition rates

Let S_t, I_t, R_t be the population sizes at time t .

$$\text{Total rate of } \{S, I\} \rightarrow \{I, I\} \quad : \quad \lambda \cdot S_t \cdot I_t.$$

$$\text{Total rate } \{R, I\} \rightarrow \{R, R\} \quad : \quad I_t \cdot R_t.$$

Hence, at time t , the probability that $\{S, I\} \rightarrow \{I, I\}$ happens before $\{R, I\} \rightarrow \{R, R\}$ is

$$\frac{\lambda S_t I_t}{\lambda S_t I_t + I_t R_t} = \frac{\lambda S_t}{\lambda S_t + R_t}.$$

Transition rates

Let S_t, I_t, R_t be the population sizes at time t .

$$\text{Total rate of } \{S, I\} \rightarrow \{I, I\} \quad : \quad \lambda \cdot S_t \cdot I_t.$$

$$\text{Total rate } \{R, I\} \rightarrow \{R, R\} \quad : \quad I_t \cdot R_t.$$

Hence, at time t , the probability that $\{S, I\} \rightarrow \{I, I\}$ happens before $\{R, I\} \rightarrow \{R, R\}$ is

$$\frac{\lambda S_t I_t}{\lambda S_t I_t + I_t R_t} = \frac{\lambda S_t}{\lambda S_t + R_t}.$$

 We are going to be able to decouple the evolutions of S and R .

COUPLING AND DECOUPLING VIA TWO YULE PROCESSES



Yule processes

Definition (Yule process)

In a Yule process $(Y(t))_{t \geq 0}$ of parameter λ , starting with one individual, each individual lives a random time distributed according to a $\text{Exp}(\lambda)$ random variable, and at its death gives birth to two individuals

Yule processes

Definition (Yule process)

In a Yule process $(Y(t))_{t \geq 0}$ of parameter λ , starting with one individual, each individual lives a random time distributed according to a $\text{Exp}(\lambda)$ random variable, and at its death gives birth to two individuals, and $Y(t)$ denotes the total number of individuals at time t .

Yule processes

Definition (Yule process)

In a Yule process $(Y(t))_{t \geq 0}$ of parameter λ , starting with one individual, each individual lives a random time distributed according to a $\text{Exp}(\lambda)$ random variable, and at its death gives birth to two individuals, and $Y(t)$ denotes the total number of individuals at time t .

 In particular, the intervals between each discontinuity are distributed according to independent $\text{Exp}(\lambda)$, $\text{Exp}(2\lambda)$, $\text{Exp}(3\lambda)$, \dots random variables.

Coupling with two Yule processes

Let $(\mathcal{R}(t))_{t \geq 0}$ be a Yule process of parameter 1, and $(\mathcal{S}_N(t))_{t \geq 0}$ a Yule process of parameter λ , time-reversed at its N -th jump.

Coupling with two Yule processes

Let $(\mathcal{R}(t))_{t \geq 0}$ be a Yule process of parameter 1, and $(\mathcal{S}_N(t))_{t \geq 0}$ a Yule process of parameter λ , time-reversed at its N -th jump.

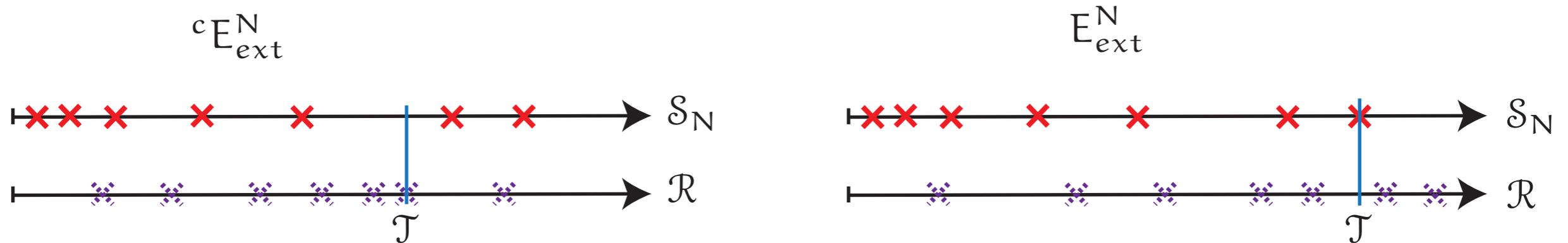


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

Coupling with two Yule processes

Let $(\mathcal{R}(t))_{t \geq 0}$ be a Yule process of parameter 1, and $(\mathcal{S}_N(t))_{t \geq 0}$ a Yule process of parameter λ , time-reversed at its N -th jump.

The prey-predator dynamics can be described by using \mathcal{R} and \mathcal{S}_N , which describe in what order the infections and recoveries happen!

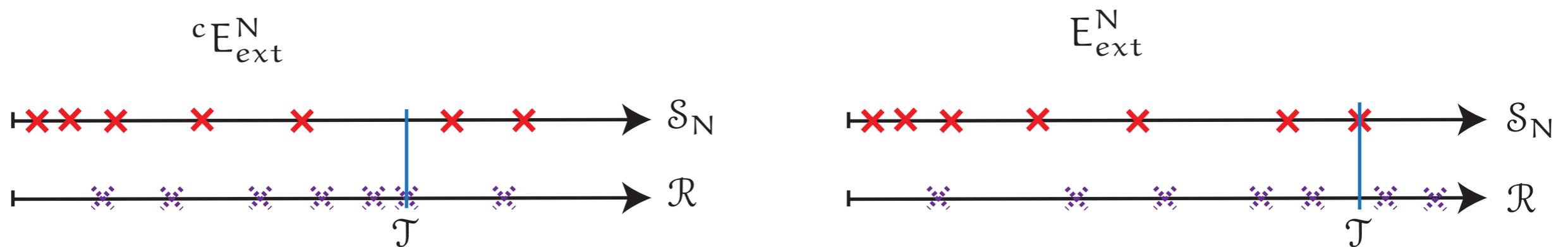


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

Coupling with two Yule processes

Let $(\mathcal{R}(t))_{t \geq 0}$ be a Yule process of parameter 1, and $(\mathcal{S}_N(t))_{t \geq 0}$ a Yule process of parameter λ , time-reversed at its N -th jump.

The prey-predator dynamics can be described by using \mathcal{R} and \mathcal{S}_N , which describe in what order the infections and recoveries happen!

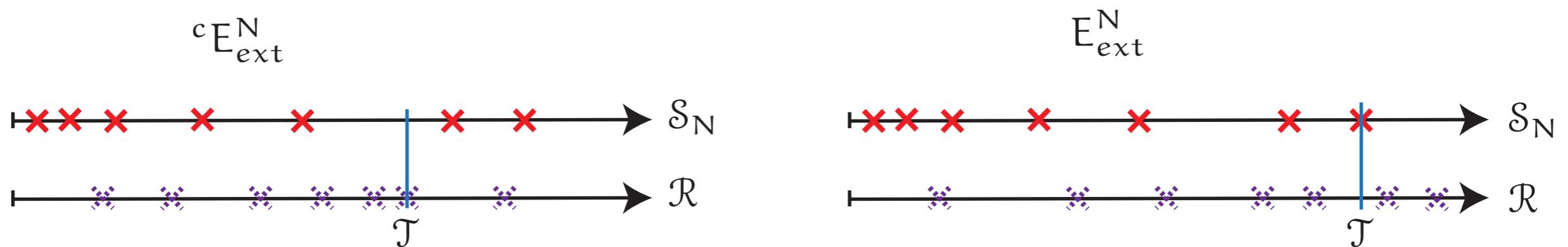


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

\mathcal{T} is the time when a type of vertices (S or I) disappears.

Coupling with two Yule processes

Let $(\mathcal{R}(t))_{t \geq 0}$ be a Yule process of parameter 1, and $(\mathcal{S}_N(t))_{t \geq 0}$ a Yule process of parameter λ , time-reversed at its N -th jump.

The prey-predator dynamics can be described by using \mathcal{R} and \mathcal{S}_N , which describe in what order the infections and recoveries happen!

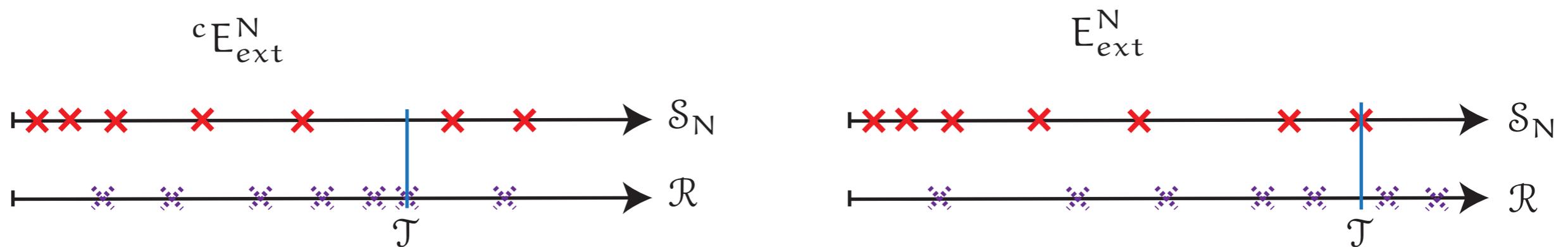


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

\mathcal{T} is the time when a type of vertices (S or I) disappears.

\mathcal{T} is the smallest between:

👉 the first moment when there are more discontinuities of \mathcal{R} than discontinuities of \mathcal{S}_N (I disappears first, ${}^c E_{\text{ext}}^N$)

👉 the N -th discontinuity of \mathcal{S}_N (S disappears first, E_{ext}^N)

IDENTIFICATION OF THE CRITICAL PARAMETER $\lambda = 1$



Notation.

Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

Notation.

Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

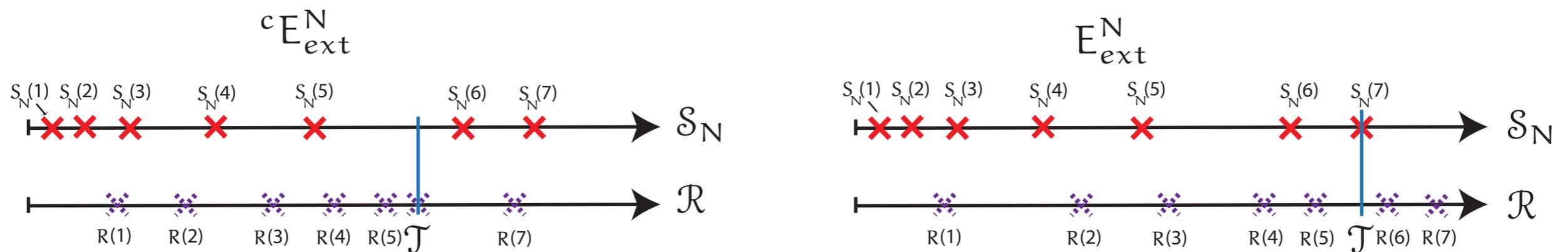


Figure: Example for $N = 7$, where the crosses represent discontinuities.

Notation.

Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

Proposition

$S_N(N)$ has the same distribution as $R(N)$ has the same distribution

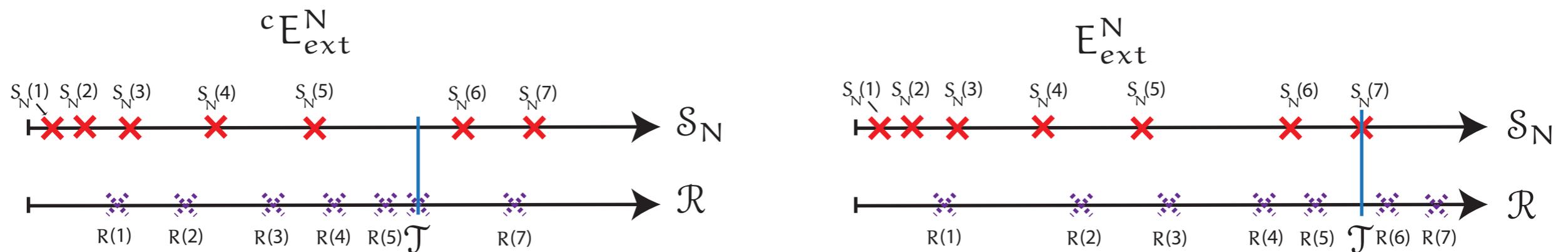


Figure: Example for $N = 7$, where the crosses represent discontinuities.

Notation.

Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

Proposition

$S_N(N)$ has the same distribution as $\text{Exp}(\lambda N) + \text{Exp}(\lambda(N-1)) + \dots + \text{Exp}(\lambda)$.
 $R(N)$ has the same distribution

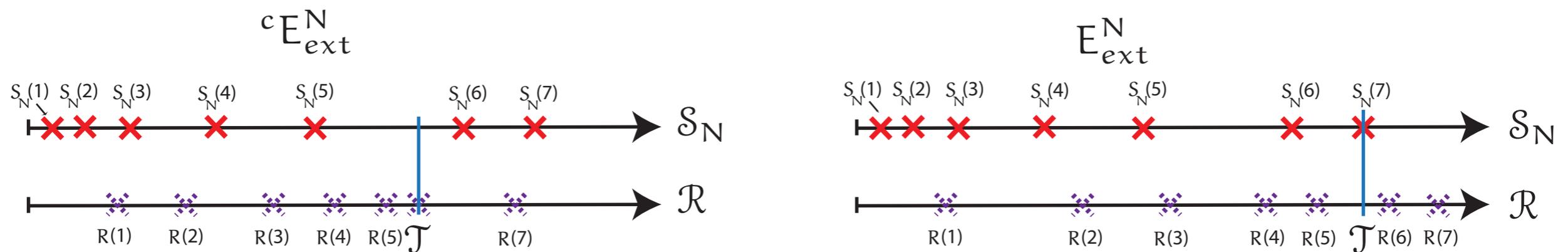


Figure: Example for $N = 7$, where the crosses represent discontinuities.

Notation.

Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

Proposition

$S_N(N)$ has the same distribution as $\text{Exp}(\lambda N) + \text{Exp}(\lambda(N-1)) + \dots + \text{Exp}(\lambda)$.
 $R(N)$ has the same distribution $\text{Exp}(1) + \text{Exp}(2) + \dots + \text{Exp}(N)$.

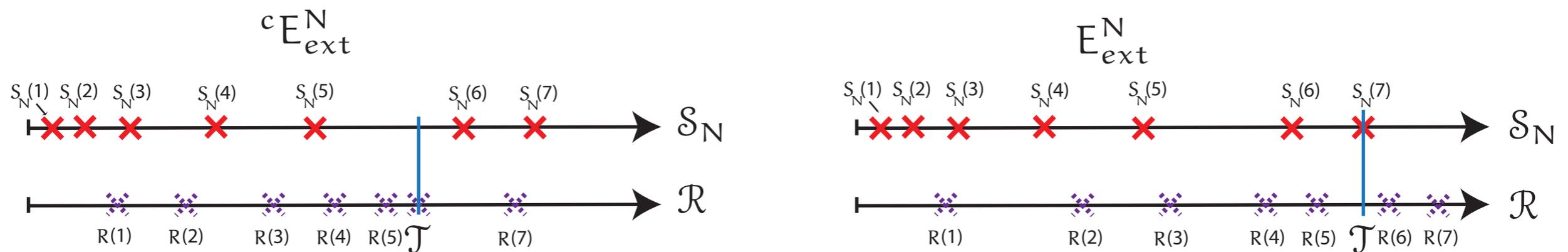


Figure: Example for $N = 7$, where the crosses represent discontinuities.

Notation.

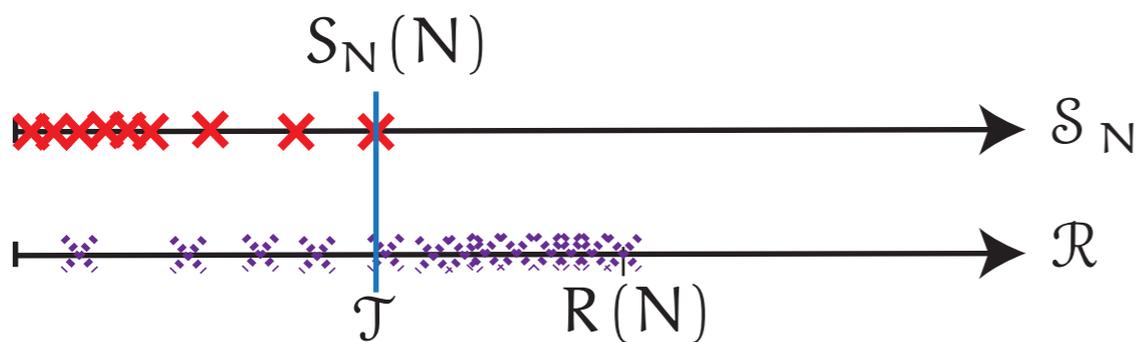
Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

Proposition

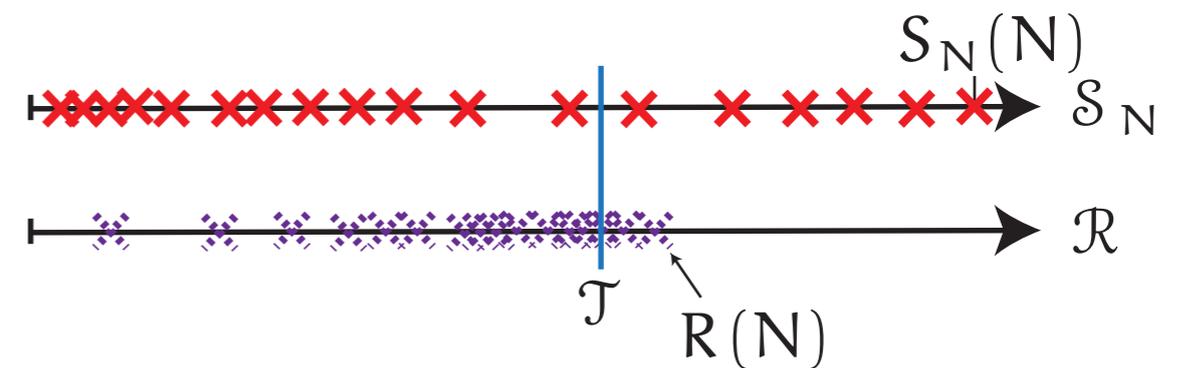
$S_N(N)$ has the same distribution as $\text{Exp}(\lambda N) + \text{Exp}(\lambda(N-1)) + \dots + \text{Exp}(\lambda)$.

$R(N)$ has the same distribution $\text{Exp}(1) + \text{Exp}(2) + \dots + \text{Exp}(N)$.

A typical situation for $\lambda > 1$:



A typical situation for $\lambda < 1$:



Notation.

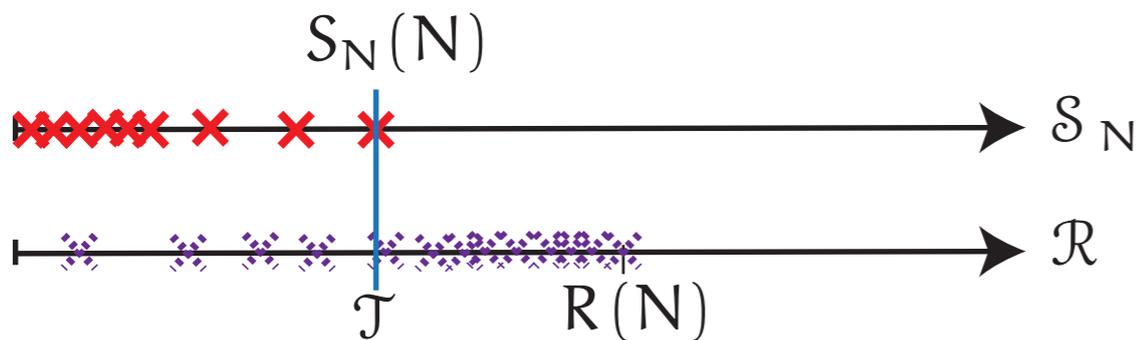
Denote by $S_N(1), S_N(2), \dots, S_N(N)$ the discontinuities \mathcal{S}_N and by $R(1), \dots, R(N)$ the discontinuities of $\mathcal{R}(t)$.

Proposition

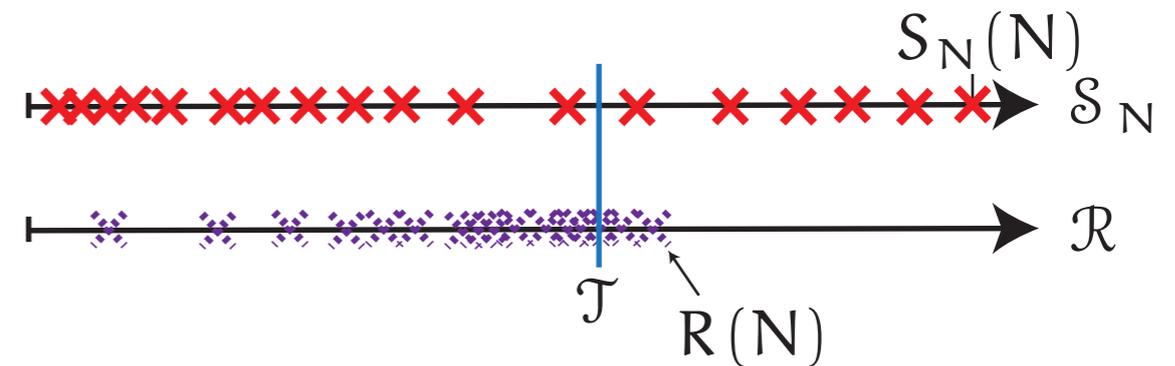
$S_N(N)$ has the same distribution as $\text{Exp}(\lambda N) + \text{Exp}(\lambda(N-1)) + \dots + \text{Exp}(\lambda)$.

$R(N)$ has the same distribution $\text{Exp}(1) + \text{Exp}(2) + \dots + \text{Exp}(N)$.

A typical situation for $\lambda > 1$:



A typical situation for $\lambda < 1$:



Hence

$$\mathbb{P}(\mathbb{E}_{\text{ext}}^N) \xrightarrow{N \rightarrow \infty} \begin{cases} 0 & \text{if } \lambda \in (0, 1) \\ \frac{1}{2} & \text{if } \lambda = 1 \\ 1 & \text{if } \lambda > 1. \end{cases}$$

STUDY OF THE FINAL STATE OF THE SYSTEM



Definition

Denote by $S^{(N)}$, $I^{(N)}$, $R^{(N)}$ the number of S , I , R vertices at the first time \mathcal{T} when a type (S or I) of vertices disappears.

Definition

Denote by $S^{(N)}$, $I^{(N)}$, $R^{(N)}$ the number of S , I , R vertices at the first time \mathcal{T} when a type (S or I) of vertices disappears.

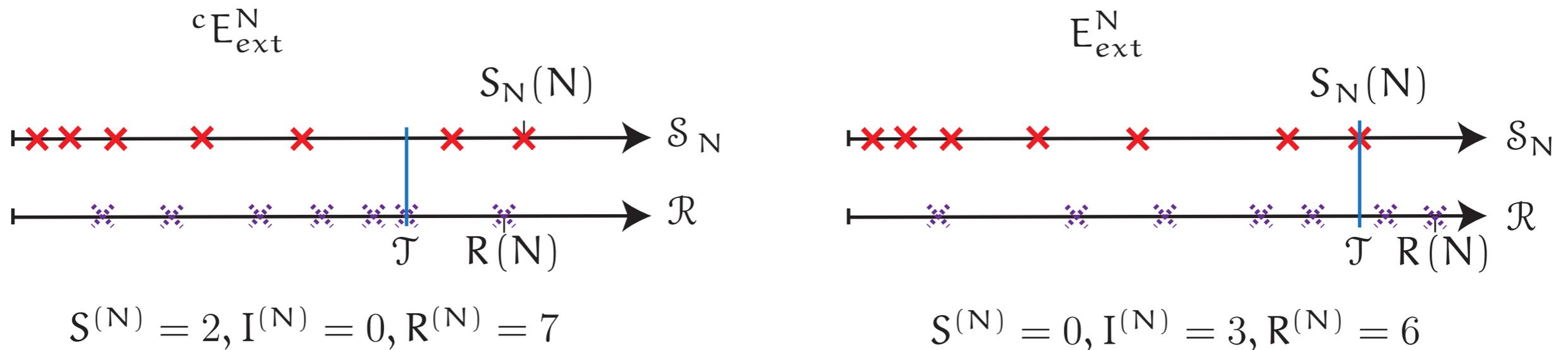


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

Definition

Denote by $S^{(N)}$, $I^{(N)}$, $R^{(N)}$ the number of S , I , R vertices at the first time \mathcal{T} when a type (S or I) of vertices disappears.

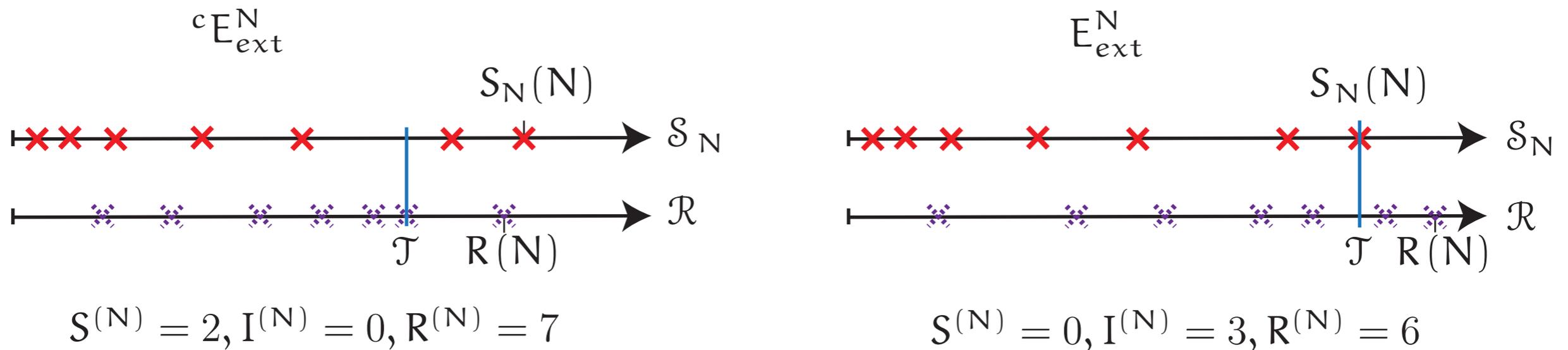


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

Question. What can be said of the asymptotic behavior of $S^{(N)}$, $I^{(N)}$, $R^{(N)}$ as $N \rightarrow \infty$?

Definition

Denote by $S^{(N)}$, $I^{(N)}$, $R^{(N)}$ the number of S , I , R vertices at the first time \mathcal{T} when a type (S or I) of vertices disappears.

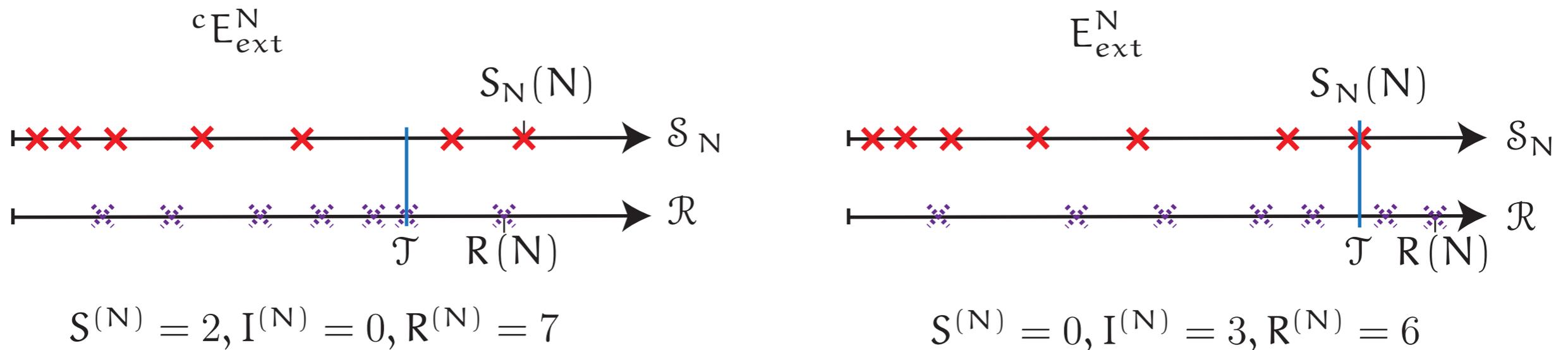


Figure: Ex. $N = 7$, where red crosses represent infections and purple ones recoveries.

Question. What can be said of the asymptotic behavior of $S^{(N)}$, $I^{(N)}$, $R^{(N)}$ as $N \rightarrow \infty$?

This should be related to the asymptotic behavior of Yule processes.

Yule processes and terminal value

Proposition

Let $(Y(t))_{t \geq 0}$ be a Yule process of parameter λ .

1) We have the convergence

$$e^{-\lambda t} Y_t \xrightarrow[t \rightarrow \infty]{\text{a.s.}} \mathcal{E},$$

where \mathcal{E} is a $\text{Exp}(1)$ random variable, called *terminal value* of Y .

Yule processes and terminal value

Proposition

Let $(Y(t))_{t \geq 0}$ be a Yule process of parameter λ .

1) We have the convergence

$$e^{-\lambda t} Y_t \xrightarrow[t \rightarrow \infty]{\text{a.s.}} \mathcal{E},$$

where \mathcal{E} is a $\text{Exp}(1)$ random variable, called *terminal value* of Y .

2) For $t \geq 0$ and $k \geq 1$, we have $\mathbb{P}(Y_t = k) = e^{-\lambda t} (1 - e^{-\lambda t})^{k-1}$.

Yule processes and terminal value

Proposition

Let $(Y(t))_{t \geq 0}$ be a Yule process of parameter λ .

1) We have the convergence

$$e^{-\lambda t} Y_t \xrightarrow[t \rightarrow \infty]{\text{a.s.}} \mathcal{E},$$

where \mathcal{E} is a $\text{Exp}(1)$ random variable, called *terminal value* of Y .

2) For $t \geq 0$ and $k \geq 1$, we have $\mathbb{P}(Y_t = k) = e^{-\lambda t} (1 - e^{-\lambda t})^{k-1}$.

Corollary

if τ_N denotes the N -th jump time of Y , then

$$\lambda \tau_N - \ln(N) \xrightarrow[N \rightarrow \infty]{\text{a.s.}} -\ln(\mathcal{E})$$

Number of susceptible individuals remaining

Theorem (K. '13).

(i) Fix $\lambda \in (0, 1)$. Then

$$\frac{S^{(N)}}{N^{1-\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(1)^\lambda.$$

(ii) Fix $\lambda = 1$.

(iii) Fix $\lambda > 1$. Then $S^{(N)}$ converges in probability towards 0 as $N \rightarrow \infty$.

Number of susceptible individuals remaining

Theorem (K. '13).

(i) Fix $\lambda \in (0, 1)$. Then

$$\frac{S^{(N)}}{N^{1-\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(1)^\lambda.$$

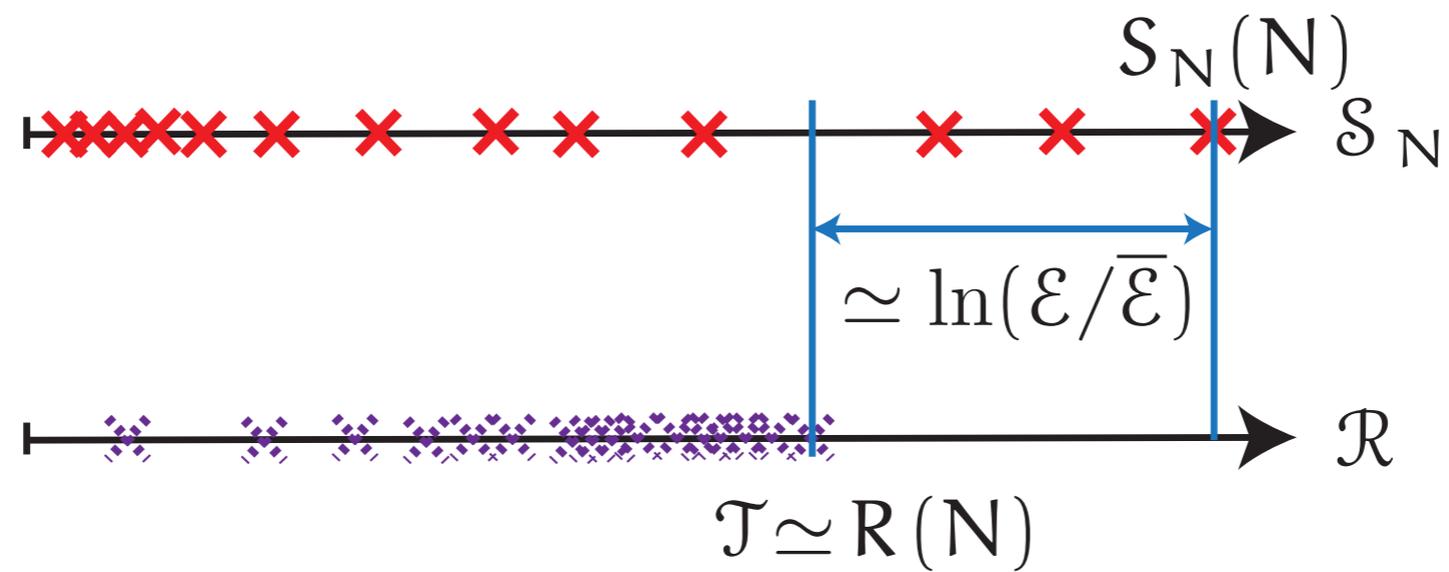
(ii) Fix $\lambda = 1$. Then for every $i \geq 0$,

$$\mathbb{P}\left(S^{(N)} = i\right) \xrightarrow[N \rightarrow \infty]{} 1/2^{i+1}.$$

(iii) Fix $\lambda > 1$. Then $S^{(N)}$ converges in probability towards 0 as $N \rightarrow \infty$.

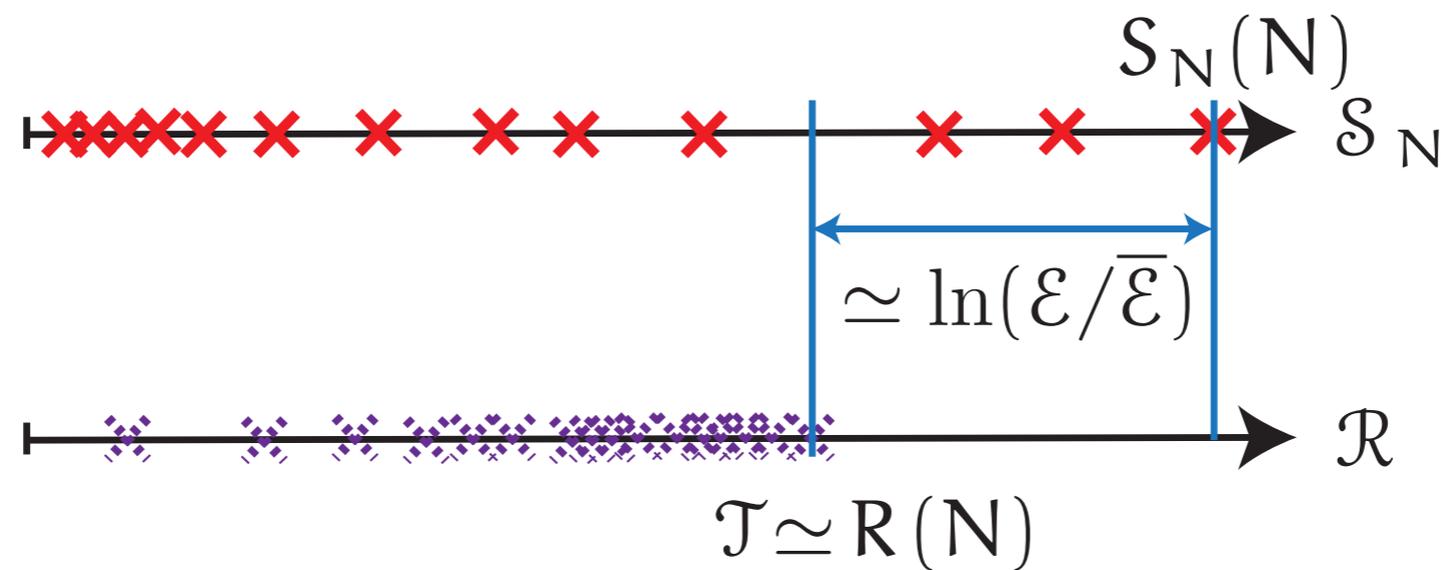
Idea of the proof: case $\lambda = 1$

On the event ${}^c E_{\text{ext}}^N$,



Idea of the proof: case $\lambda = 1$

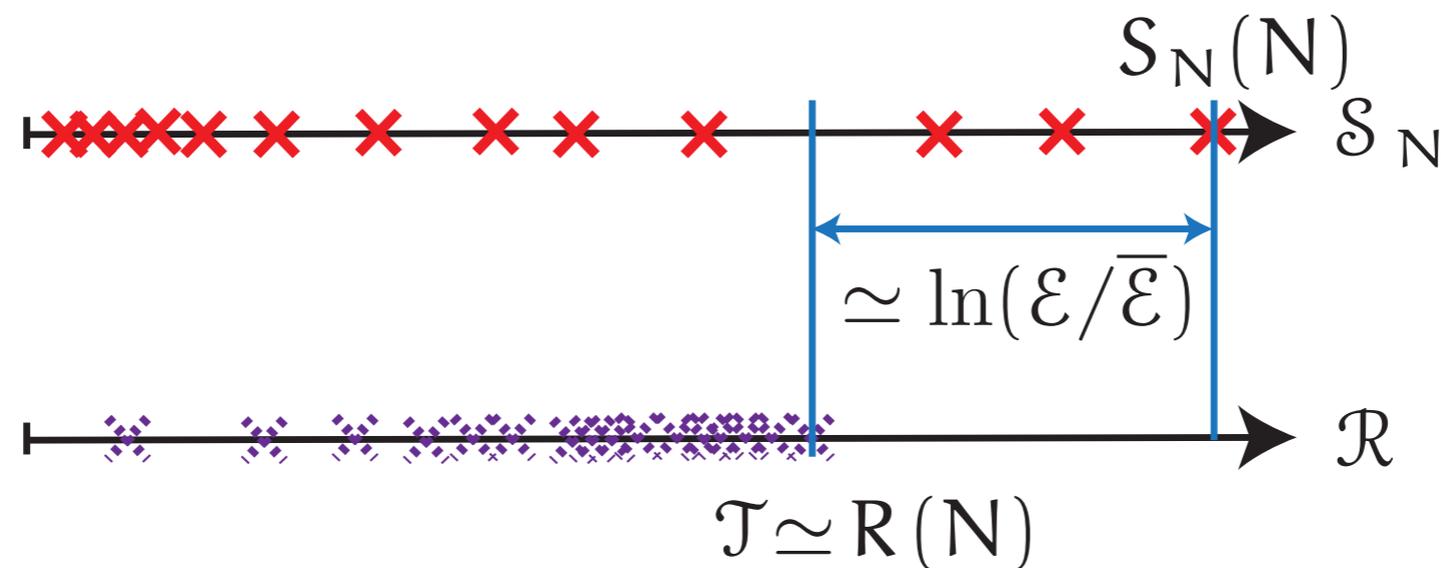
On the event ${}^c E_{\text{ext}}^N$,



Let $\bar{\mathcal{E}}$ be the terminal value of the Yule process associated with S_N , and \mathcal{E} is the terminal value of \mathcal{R} .

Idea of the proof: case $\lambda = 1$

On the event ${}^c E_{\text{ext}}^N$,

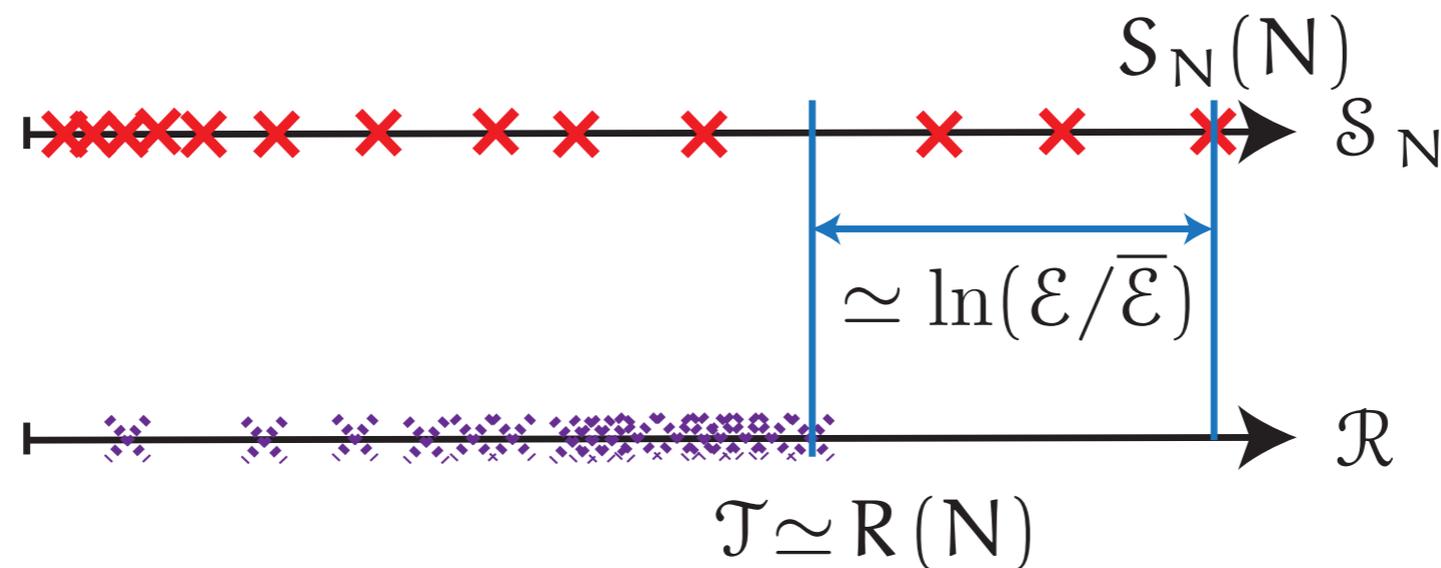


Let $\bar{\mathcal{E}}$ be the terminal value of the Yule process associated with S_N , and \mathcal{E} is the terminal value of \mathcal{R} .

We have $S_N(N) \simeq \ln(N) - \ln(\bar{\mathcal{E}})$, $R(N) \simeq \ln(N) - \ln(\mathcal{E})$

Idea of the proof: case $\lambda = 1$

On the event ${}^c E_{\text{ext}}^N$,

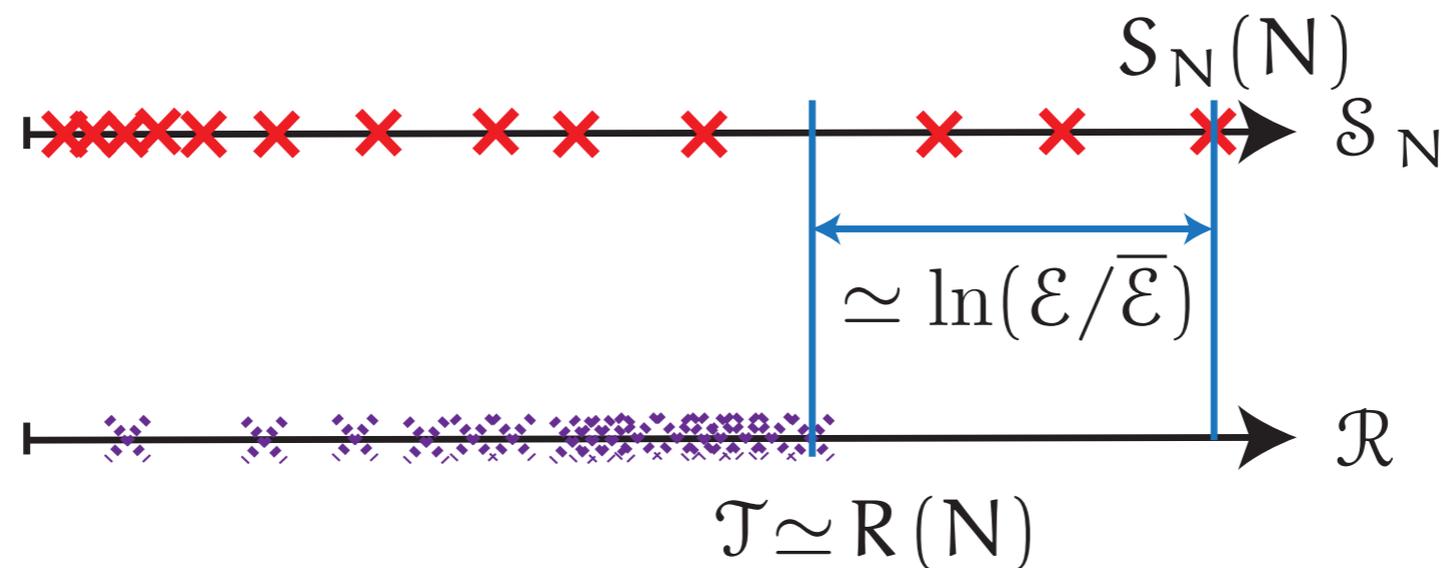


Let $\bar{\mathcal{E}}$ be the terminal value of the Yule process associated with S_N , and \mathcal{E} is the terminal value of \mathcal{R} .

We have $S_N(N) \simeq \ln(N) - \ln(\bar{\mathcal{E}})$, $R(N) \simeq \ln(N) - \ln(\mathcal{E})$, with $\mathcal{E}/\bar{\mathcal{E}} > 1$.

Idea of the proof: case $\lambda = 1$

On the event ${}^c E_{\text{ext}}^N$,

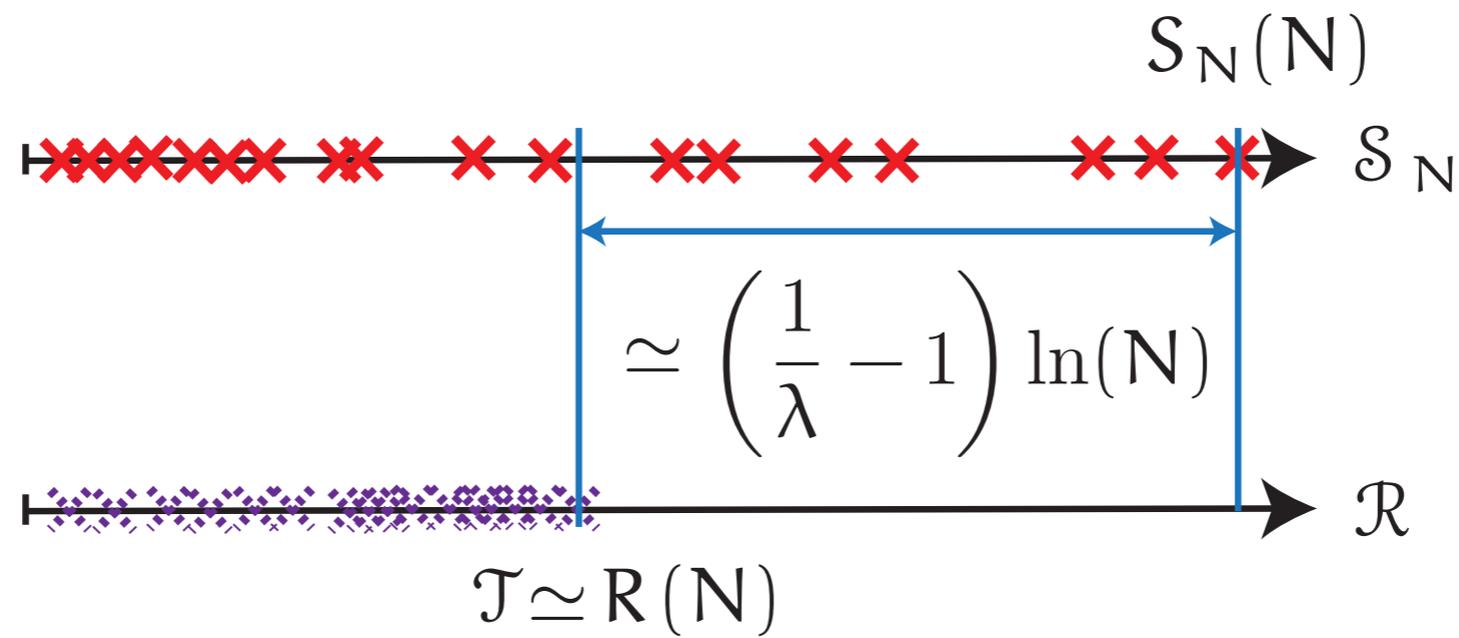


Let $\bar{\mathcal{E}}$ be the terminal value of the Yule process associated with S_N , and \mathcal{E} is the terminal value of \mathcal{R} .

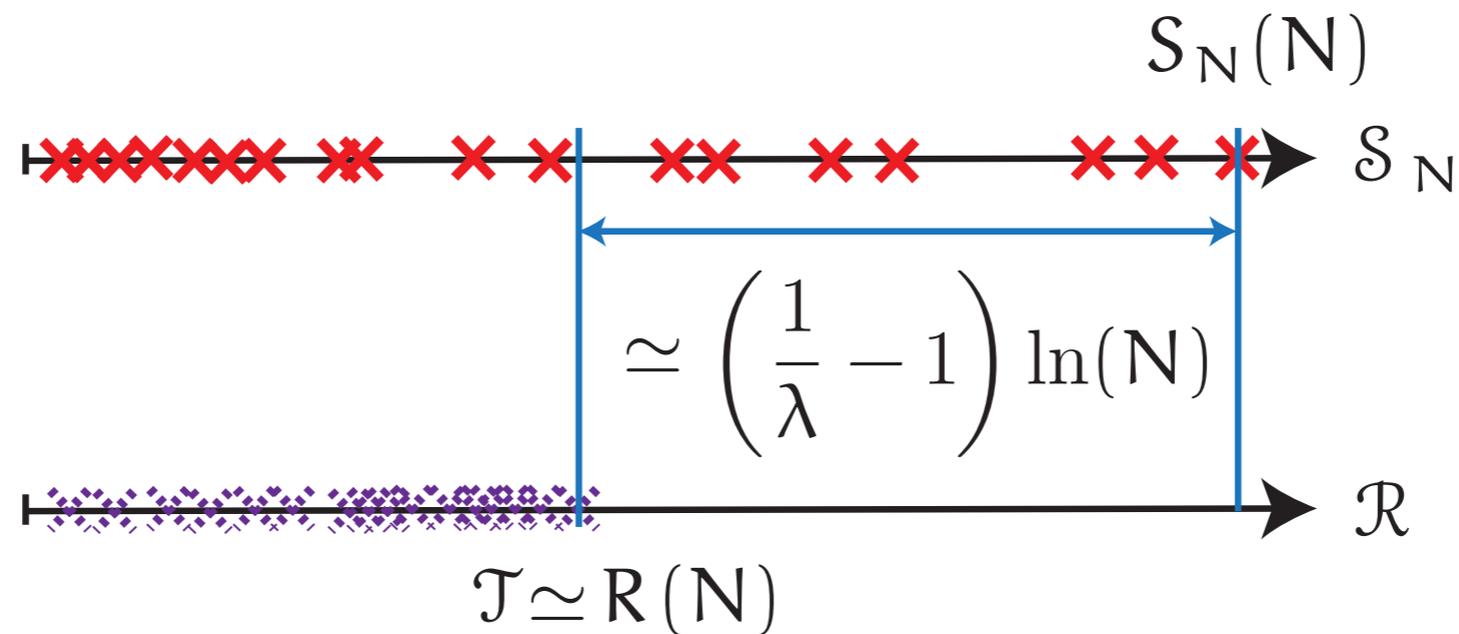
We have $S_N(N) \simeq \ln(N) - \ln(\bar{\mathcal{E}})$, $\mathcal{R}(N) \simeq \ln(N) - \ln(\mathcal{E})$, with $\mathcal{E}/\bar{\mathcal{E}} > 1$.

Thus, $S^{(N)}$ \simeq value of a Yule process of parameter λ at time $\ln(\mathcal{E}/\bar{\mathcal{E}})$, conditionnally on $\mathcal{E}/\bar{\mathcal{E}} > 1$.

Idea of proof: case $\lambda \in (0, 1)$

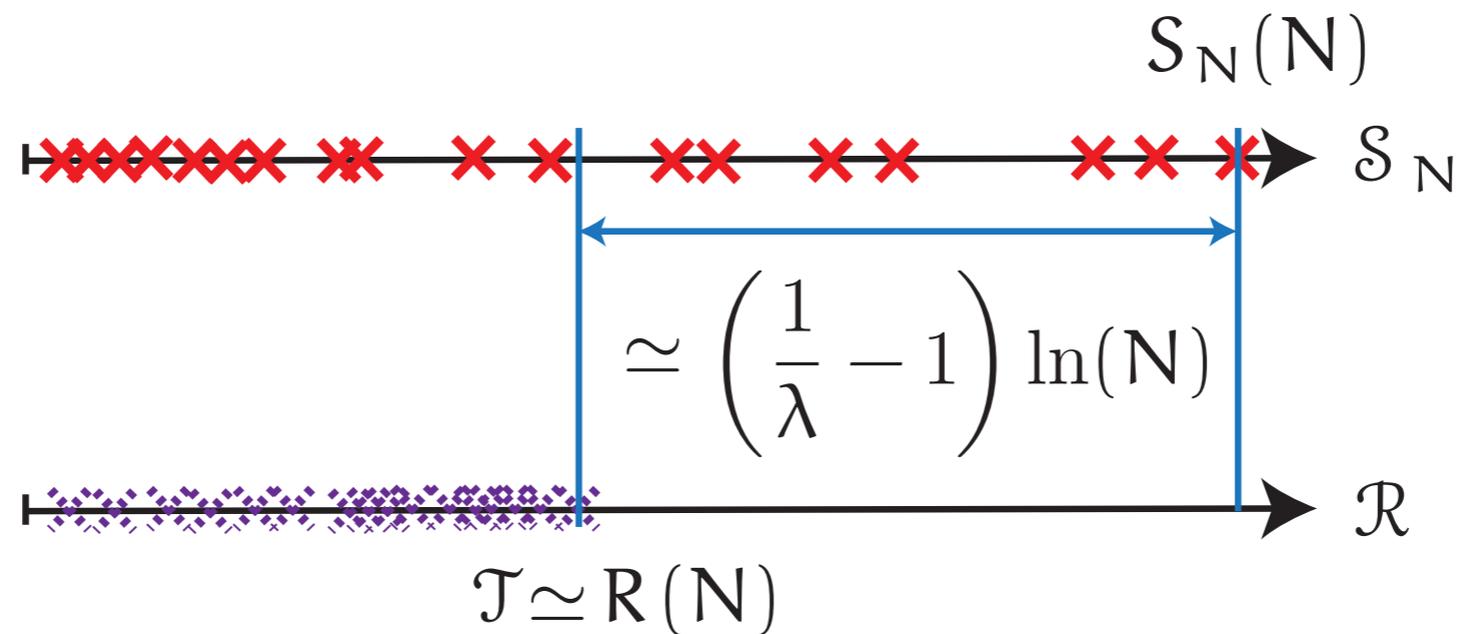


Idea of proof: case $\lambda \in (0, 1)$



Recall that $\bar{\mathcal{E}}$ is the terminal value of the Yule process associated with \mathcal{S}_N , and \mathcal{E} is the terminal value of \mathcal{R} .

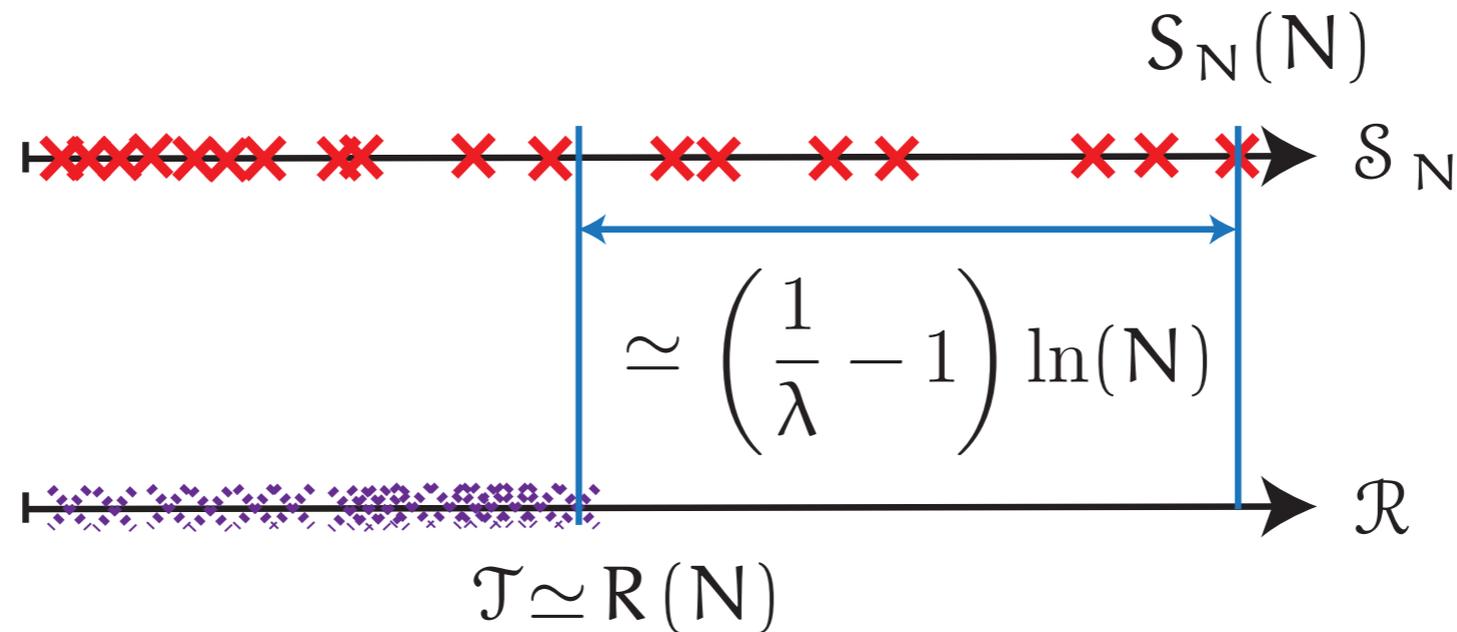
Idea of proof: case $\lambda \in (0, 1)$



Recall that $\bar{\mathcal{E}}$ is the terminal value of the Yule process associated with S_N , and \mathcal{E} is the terminal value of \mathcal{R} .

We have $S_N(N) \simeq \frac{1}{\lambda} (\ln(N) - \ln(\bar{\mathcal{E}}))$, $\mathcal{R}(N) \simeq \ln(N) - \ln(\mathcal{E})$.

Idea of proof: case $\lambda \in (0, 1)$

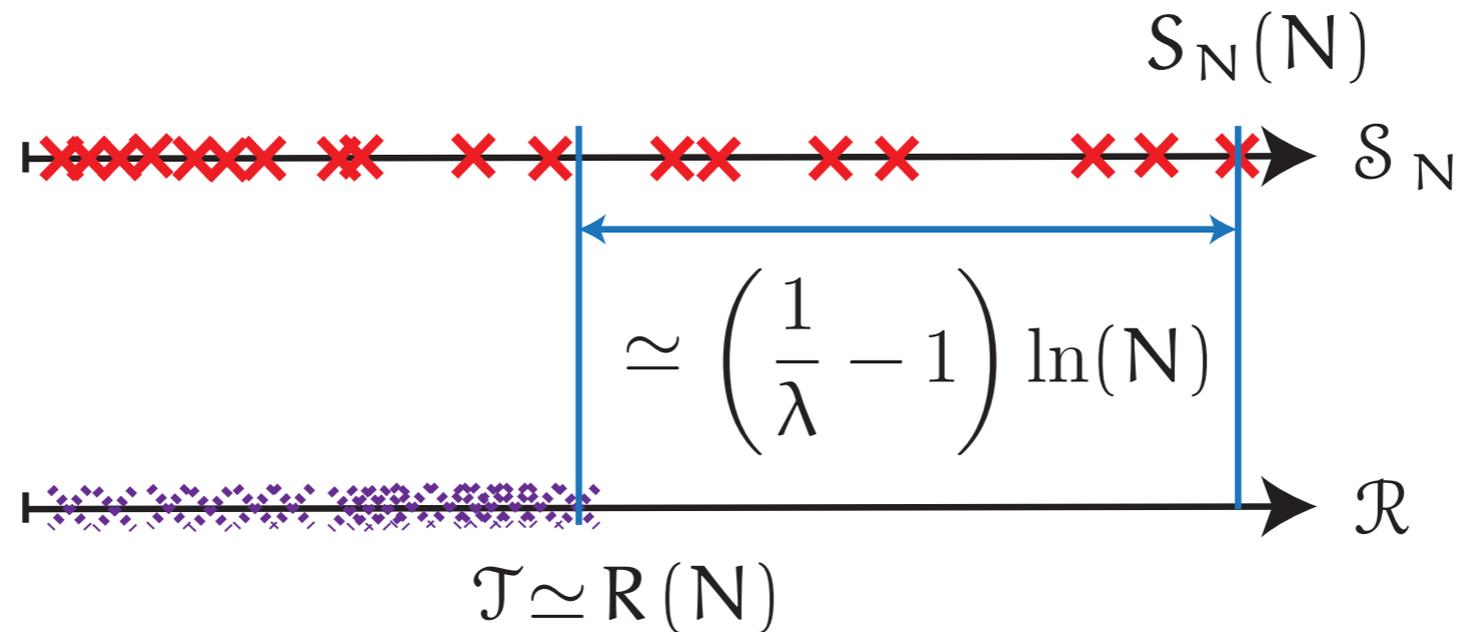


Recall that $\bar{\mathcal{E}}$ is the terminal value of the Yule process associated with \mathcal{S}_N , and \mathcal{E} is the terminal value of \mathcal{R} .

We have $S_N(N) \simeq \frac{1}{\lambda} (\ln(N) - \ln(\bar{\mathcal{E}}))$, $\mathcal{R}(N) \simeq \ln(N) - \ln(\mathcal{E})$.

Thus, $S^{(N)}$ \simeq value of a Yule process of parameter λ at time $(1/\lambda - 1) \ln(N)$.

Idea of proof: case $\lambda \in (0, 1)$



Recall that $\bar{\mathcal{E}}$ is the terminal value of the Yule process associated with \mathcal{S}_N , and \mathcal{E} is the terminal value of \mathcal{R} .

We have $S_N(N) \simeq \frac{1}{\lambda} (\ln(N) - \ln(\bar{\mathcal{E}}))$, $\mathcal{R}(N) \simeq \ln(N) - \ln(\mathcal{E})$.

Thus, $S^{(N)} \simeq$ value of a Yule process of parameter λ at time $(1/\lambda - 1) \ln(N)$. Which is of order $e^{\lambda(1/\lambda - 1) \ln(N)} = N^{1-\lambda}$.

Number of recovered individuals remaining

Theorem (K. '13).

(i) Fix $\lambda \in (0, 1)$. Then

$$\frac{N - R^{(N)}}{N^{1-\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(1)^\lambda.$$

(ii) Fix $\lambda = 1$.

(iii) Fix $\lambda > 1$. Then

Number of recovered individuals remaining

Theorem (K. '13).

(i) Fix $\lambda \in (0, 1)$. Then

$$\frac{N - \mathcal{R}^{(N)}}{N^{1-\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(1)^\lambda.$$

(ii) Fix $\lambda = 1$. Then

$$\frac{\mathcal{R}^{(N)}}{N} \xrightarrow[N \rightarrow \infty]{(d)} \frac{1}{2} \delta_1 +$$

(iii) Fix $\lambda > 1$. Then

Number of recovered individuals remaining

Theorem (K. '13).

(i) Fix $\lambda \in (0, 1)$. Then

$$\frac{N - \mathcal{R}^{(N)}}{N^{1-\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(1)^\lambda.$$

(ii) Fix $\lambda = 1$. Then

$$\frac{\mathcal{R}^{(N)}}{N} \xrightarrow[N \rightarrow \infty]{(d)} \frac{1}{2} \delta_1 + \frac{1}{(1+x)^2} \mathbb{1}_{[0,1]}(x) dx,$$

(iii) Fix $\lambda > 1$. Then

Number of recovered individuals remaining

Theorem (K. '13).

(i) Fix $\lambda \in (0, 1)$. Then

$$\frac{N - \mathcal{R}^{(N)}}{N^{1-\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(1)^\lambda.$$

(ii) Fix $\lambda = 1$. Then

$$\frac{\mathcal{R}^{(N)}}{N} \xrightarrow[N \rightarrow \infty]{(d)} \frac{1}{2} \delta_1 + \frac{1}{(1+x)^2} \mathbb{1}_{[0,1]}(x) dx,$$

(iii) Fix $\lambda > 1$. Then

$$\frac{\mathcal{R}^{(N)}}{N^{1/\lambda}} \xrightarrow[N \rightarrow \infty]{(d)} \text{Exp}(\text{Exp}(1)^{1/\lambda}).$$

Calculations involving Yule processes

Key idea: Kendall's representaton of Yule processes.

Calculations involving Yule processes

Key idea: Kendall's representation of Yule processes.

Theorem (Kendall '66)

Let $(\mathcal{P}_t)_{t \geq 0}$ be a Poisson process of parameter 1 starting from 0, and \mathcal{E} be an exponential random variable of parameter 1. Then

$$t \mapsto \mathcal{P}_{\mathcal{E}(e^{\lambda t} - 1)} + 1$$

is a Yule process of parameter λ with terminal value \mathcal{E} .

Calculations involving Yule processes

Key idea: Kendall's representation of Yule processes.

Theorem (Kendall '66)

Let $(\mathcal{P}_t)_{t \geq 0}$ be a Poisson process of parameter 1 starting from 0, and ε be an exponential random variable of parameter 1. Then

$$t \mapsto \mathcal{P}_{\varepsilon(e^{\lambda t} - 1)} + 1$$

is a Yule process of parameter λ with terminal value ε .

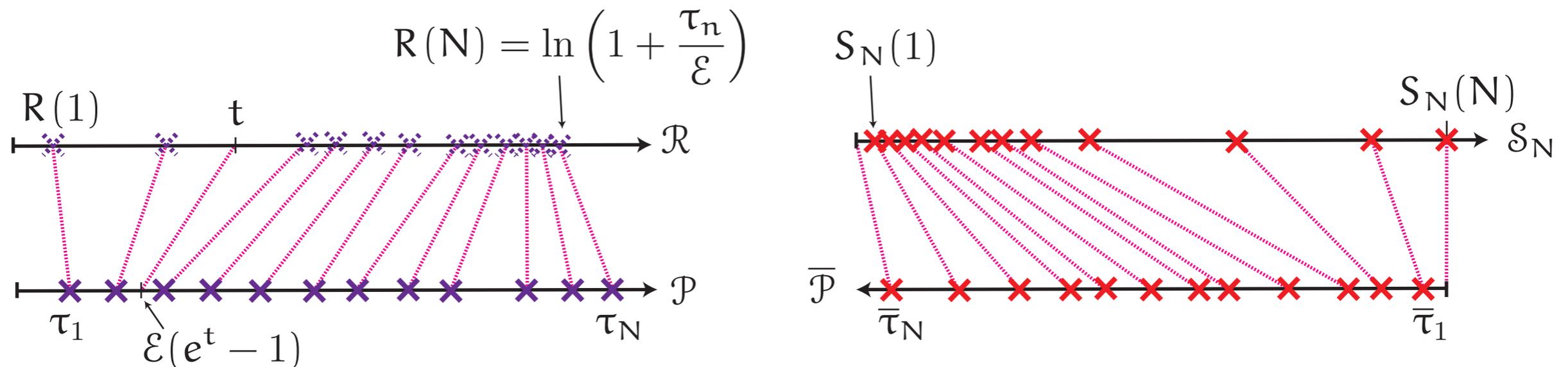


Figure: Illustration of the coupling of Yule processes with Poisson processes

Calculations involving Yule processes

Key idea: Kendall's representation of Yule processes.

Theorem (Kendall '66)

Let $(\mathcal{P}_t)_{t \geq 0}$ be a Poisson process of parameter 1 starting from 0, and \mathcal{E} be an exponential random variable of parameter 1. Then

$$t \mapsto \mathcal{P}_{\mathcal{E}(e^{\lambda t} - 1)} + 1$$

is a Yule process of parameter λ with terminal value \mathcal{E} .

This allows to calculate explicitly the limiting laws in the previous theorems

Calculations involving Yule processes

Key idea: Kendall's representation of Yule processes.

Theorem (Kendall '66)

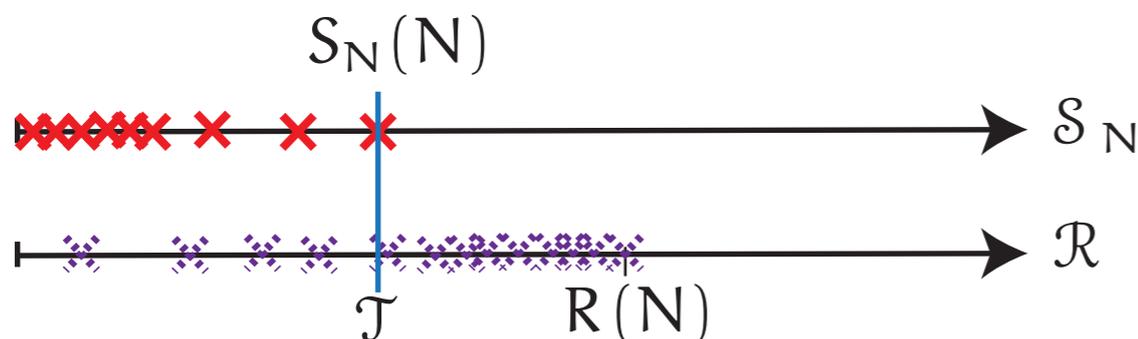
Let $(\mathcal{P}_t)_{t \geq 0}$ be a Poisson process of parameter 1 starting from 0, and \mathcal{E} be an exponential random variable of parameter 1. Then

$$t \mapsto \mathcal{P}_{\mathcal{E}(e^{\lambda t} - 1)} + 1$$

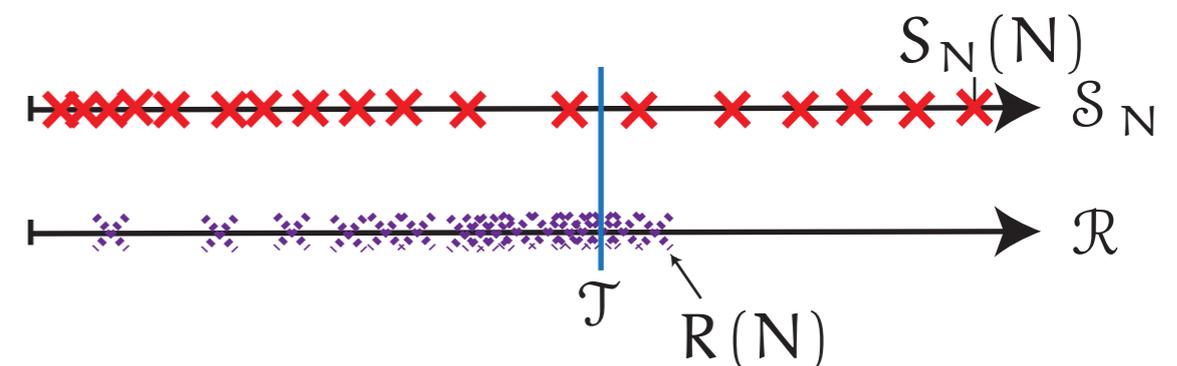
is a Yule process of parameter λ with terminal value \mathcal{E} .

This allows to calculate explicitly the limiting laws in the previous theorems, and to justify the approximation:

A typical situation for $\lambda > 1$:



A typical situation for $\lambda < 1$:



I. TEST YOUR INTUITION!

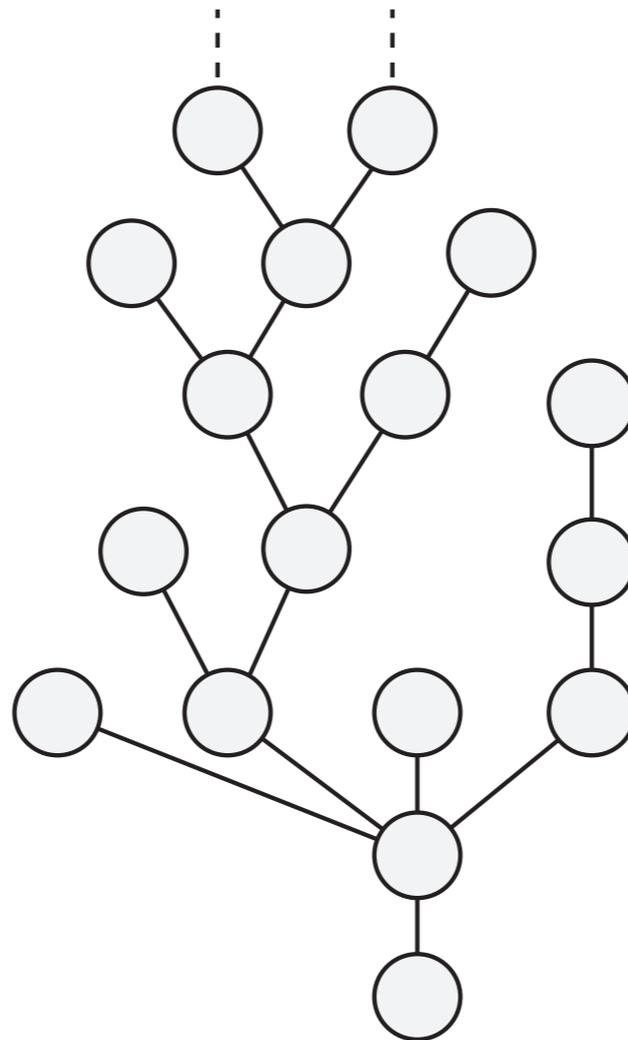
II. PREYS & PREDATORS ON A COMPLETE GRAPH

III. PREYS & PREDATORS ON AN INFINITE TREE



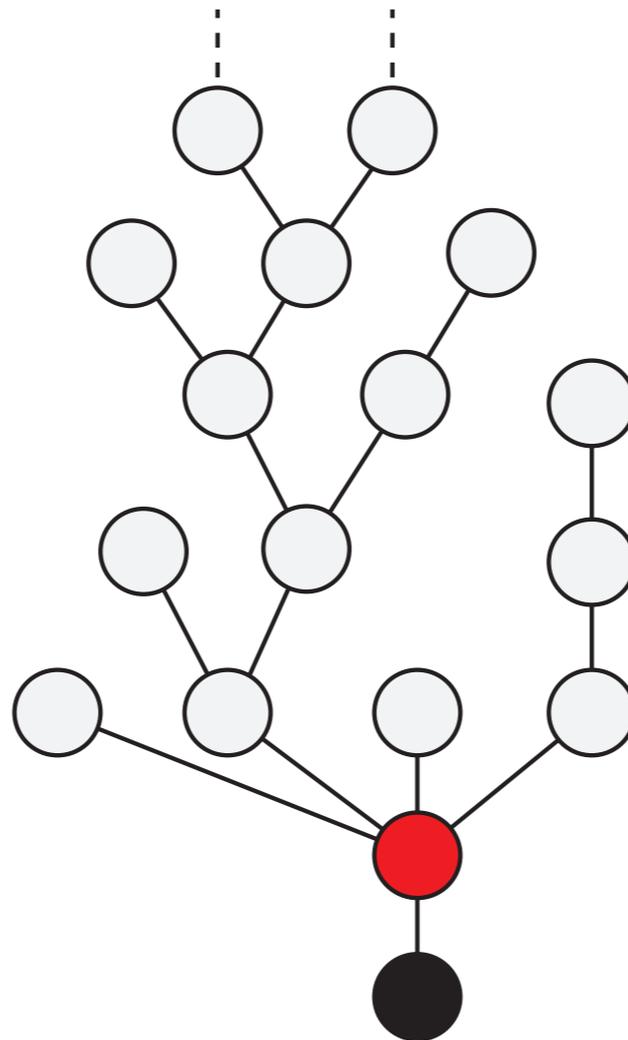
Prey-predators on trees

Let T be a rooted tree, and \hat{T} be the tree obtained by adding a parent to the root of T .



Prey-predators on trees

Let T be a rooted tree, and \hat{T} be the tree obtained by adding a parent to the root of T . Start the prey-predator process with one predator at the root of \hat{T} and a prey at the root of T .



What is the probability $p_T(\lambda)$ that the preys survive indefinitely?

Prey-predators on Galton–Watson trees

Let ν be a probability measure on \mathbb{Z}_+ . Set $d := \sum_{i \geq 0} i\nu(i)$ and assume that $d > 1$. Let \mathcal{T} be a Galton–Watson tree with offspring distribution ν .

Prey-predators on Galton–Watson trees

Let ν be a probability measure on \mathbb{Z}_+ . Set $d := \sum_{i \geq 0} i\nu(i)$ and assume that $d > 1$. Let \mathcal{T} be a Galton–Watson tree with offspring distribution ν .

Theorem (Kordzakhia '05)

If \mathcal{T} is an infinite d -ary tree, and

$$\lambda_c := 2d - 1 - 2\sqrt{d(d-1)},$$

then $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda < \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Prey-predators on Galton–Watson trees

Let ν be a probability measure on \mathbb{Z}_+ . Set $d := \sum_{i \geq 0} i\nu(i)$ and assume that $d > 1$. Let \mathcal{T} be a Galton–Watson tree with offspring distribution ν .

Theorem (Kordzakhia '05)

If \mathcal{T} is an infinite d -ary tree, and

$$\lambda_c := 2d - 1 - 2\sqrt{d(d-1)},$$

then $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda < \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Theorem (Bordenave '12)

Almost surely, we have $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda \leq \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Prey-predators on Galton–Watson trees

Let ν be a probability measure on \mathbb{Z}_+ . Set $d := \sum_{i \geq 0} i\nu(i)$ and assume that $d > 1$. Let \mathcal{T} be a Galton–Watson tree with offspring distribution ν .

Theorem (Kordzakhia '05)

If \mathcal{T} is an infinite d -ary tree, and

$$\lambda_c := 2d - 1 - 2\sqrt{d(d-1)},$$

then $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda < \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Theorem (Bordenave '12)

Almost surely, we have $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda \leq \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Denote by Z the total number of **Infected** individuals.

Prey-predators on Galton–Watson trees

Let ν be a probability measure on \mathbb{Z}_+ . Set $d := \sum_{i \geq 0} i\nu(i)$ and assume that $d > 1$. Let \mathcal{T} be a Galton–Watson tree with offspring distribution ν .

Theorem (Kordzakhia '05)

If \mathcal{T} is an infinite d -ary tree, and

$$\lambda_c := 2d - 1 - 2\sqrt{d(d-1)},$$

then $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda < \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Theorem (Bordenave '12)

Almost surely, we have $p_{\mathcal{T}}(\lambda) = 0$ for $\lambda \leq \lambda_c$ and $p_{\mathcal{T}}(\lambda) > 0$ for $\lambda > \lambda_c$.

Denote by Z the total number of **Infected** individuals.

Theorem (Bordenave '12)

If $\lambda < \lambda_c$, we have (under an integrability assumption on ν)

$$\sup\{u \geq 1; \mathbb{E}[Z^u] < \infty\} = \frac{(1 - \lambda + \sqrt{\lambda^2 - 2\lambda(2d - 1) + 1})^2}{4(d - 1)\lambda}$$

Tail of the number of Infected individuals

Theorem (K. '13).

(i) Assume that $\lambda = \lambda_c$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim}$$

(ii) Assume that $\lambda \in (0, \lambda_c)$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim}$$

Tail of the number of Infected individuals

Theorem (K. '13).

(i) Assume that $\lambda = \lambda_c$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} \left(1 + \sqrt{\frac{d}{d-1}}\right) \cdot \frac{1}{n(\ln(n))^2}.$$

(ii) Assume that $\lambda \in (0, \lambda_c)$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim}$$

Tail of the number of Infected individuals

Theorem (K. '13).

(i) Assume that $\lambda = \lambda_c$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} \left(1 + \sqrt{\frac{d}{d-1}}\right) \cdot \frac{1}{n(\ln(n))^2}.$$

(ii) Assume that $\lambda \in (0, \lambda_c)$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} C(\lambda, d) \cdot n^{-\frac{(1-\lambda + \sqrt{\lambda^2 - 2\lambda(2d-1)+1})^2}{4(d-1)\lambda}}.$$

Tail of the number of Infected individuals

Theorem (K. '13).

(i) Assume that $\lambda = \lambda_c$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} \left(1 + \sqrt{\frac{d}{d-1}}\right) \cdot \frac{1}{n(\ln(n))^2}.$$

(ii) Assume that $\lambda \in (0, \lambda_c)$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} C(\lambda, d) \cdot n^{-\frac{(1-\lambda + \sqrt{\lambda^2 - 2\lambda(2d-1)+1})^2}{4(d-1)\lambda}}.$$

For $\lambda = \lambda_c$, we have $\mathbb{E}[Z] < \infty$, but $\mathbb{E}[Z \ln(Z)] = \infty$.

Tail of the number of Infected individuals

Theorem (K. '13).

(i) Assume that $\lambda = \lambda_c$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} \left(1 + \sqrt{\frac{d}{d-1}}\right) \cdot \frac{1}{n(\ln(n))^2}.$$

(ii) Assume that $\lambda \in (0, \lambda_c)$. Then

$$\mathbb{P}(Z > n) \underset{n \rightarrow \infty}{\sim} C(\lambda, d) \cdot n^{-\frac{(1-\lambda + \sqrt{\lambda^2 - 2\lambda(2d-1) + 1})^2}{4(d-1)\lambda}}.$$

For $\lambda = \lambda_c$, we have $\mathbb{E}[Z] < \infty$, but $\mathbb{E}[Z \ln(Z)] = \infty$.

 Idea: explicit coupling with a branching random walk killed at the origin, and use results of Aïdékon, Hu & Zindy.

Coupling with a branching random walk

Let V be the branching random walk produced with the point process

$$\mathcal{L} = \sum_{i=1}^{\mathcal{U}} \delta_{\{\mathcal{E} - \text{Exp}_i(\lambda)\}},$$

starting from 0, where \mathcal{U} is a r.v distributed as ν , where \mathcal{E} is an independent $\text{Exp}(1)$ r.v and $(\text{Exp}_i(\lambda))_{i \geq 1}$ are independent i.i.d. $\text{Exp}(\lambda)$.

Coupling with a branching random walk

Let V be the branching random walk produced with the point process

$$\mathcal{L} = \sum_{i=1}^{\mathcal{U}} \delta_{\{\mathcal{E} - \text{Exp}_i(\lambda)\}},$$

starting from 0, where \mathcal{U} is a r.v distributed as ν , where \mathcal{E} is an independent $\text{Exp}(1)$ r.v and $(\text{Exp}_i(\lambda))_{i \geq 1}$ are independent i.i.d. $\text{Exp}(\lambda)$.

Kill V at 0, by only considering $\{u \in \mathcal{T}; V(v) \geq 0, \forall v \in [[\emptyset, u]]\}$.

Coupling with a branching random walk

Let V be the branching random walk produced with the point process

$$\mathcal{L} = \sum_{i=1}^{\mathcal{U}} \delta_{\{\mathcal{E} - \text{Exp}_i(\lambda)\}},$$

starting from 0, where \mathcal{U} is a r.v distributed as ν , where \mathcal{E} is an independent $\text{Exp}(1)$ r.v and $(\text{Exp}_i(\lambda))_{i \geq 1}$ are independent i.i.d. $\text{Exp}(\lambda)$.

Kill V at 0, by only considering $\{\mathfrak{u} \in \mathcal{T}; V(\nu) \geq 0, \forall \nu \in [[\emptyset, \mathfrak{u}]]\}$.

Proposition.

The number Z of infected individuals has the same distribution as

$$\#\{\mathfrak{u} \in \mathcal{T}; V(\nu) \geq 0, \forall \nu \in [[\emptyset, \mathfrak{u}]]\}.$$